#### UNIVERSITY FOR DEVELOPMENT STUDIES, TAMALE

## SURVIVAL AND PROGNOSTIC FACTORS OF HIV/AIDS, TB AND CO-INFECTION IN PRU DISTRICT

BY

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DEGREE IN APPLIED STATISTICS

#### **DECLARATION**

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I hereby declare that this thesis is the result of my own original work and that no part of it has been presented for another degree in this university or elsewhere.

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#### **ABSTRACT**

Survival analysis is a method for analysing the occurrence of a given event. This research seeks to evaluate the survival of HIV/AIDS, TB and co-infected patients and to identify the major prognostic factors that influence their survival. In this study, survival data for HIV/AIDS, TB and HIV/TB coinfection patients were obtained from St. Mathias Hospital in Yeji, Pru District of Brong Ahafo Region of Ghana. The data was fitted using both the Cox model and accelerated failure time model. The AFT (Gamma) was the best model for HIV/AIDS and HIV/TB co-infection survival data, However, the Cox proportional hazard model was best for the TB data based on the AIC and BIC values. The study revealed that none of the covariates significantly interact at 10% significance level. The diagnostic checks on the Cox-Snell residual plot of the gamma model shows that it has the best predictive power because it is closer to the bisector. Cox model revealed that the proportionality assumption was satisfied. The martingale residual plot of the continuous covariates indicate that for each of the covariates, the plot do not show trend and the resulting smoothed plots (LOESS) are approximately horizontal straight lines. This confirms that the martingale residual plots have a linear relationship with the survival time. Hence, the model is adequate. The HIV/TB co-infection patients experienced the worse survival rate. The study deduced that Weight significantly determines the patient's survival among all the three categories. Therefore health authorities should be very cautious and pay much attention to patients who weighed below the minimum weights.

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## **DEDICATION**

This work is dedicated to my beloved mother Janet Akosua Yaka and father Nawumbeni Dabanyi.

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#### LIST OF ACRONYMS

3TC Lamivudine

AFTM Accelerated Failure Time Model

AIC Akaike Information Criterion

AIDS Acquired Immunodeficiency Syndrome

ART Antiretroviral therapy

ARV Antiretroviral

AZT Zidovudine

BIC Bayesian Information Criterion

CBV Combivir

CD4 Cluster of Differentiation Four

DOTS Directly Observed Treatment Short-Course

EFV Efavirenz

FANTA Food and Nutrition Technical Assistance

GDHS Ghana Demographic and Health Survey

GHS Ghana Health Service

HAART Highly Active Antiretroviral Therapy

HIV Human Immunodeficiency Virus

HR Hazard Ratio

HRZ Isoniazid Rifampicin Pyrazinamide

HRZE Isoniazid Rifampicin Pyrazinamide Ethambotol

HRZES Isoniazid Rifampicin Pyrazinamide Ethambotol Steptomymide

Kg Kilogram

KM Kaplan-Meier Estimator

LR Likelihood Ratio

Maximum Likelihood Estimate/Estimator

MOH

MLE

Ministry of Health

**NACP** 

National AIDS Control Programmes

**NFHL** 

Nutrition for Healthy Living

NTP

National Tuberculosis Programme

NVP

Nevirapine

PH

**Proportional Hazard** 

**PLWHA** 

People Living with HIV/AIDS

TB

Tuberculosis

TR

Time Ratio

WHO

World Health Organization



#### CHAPTER ONE

#### INTRODUCTION

#### 1.1 Background of the study

The complex relationship between Human Immunodeficiency Virus (HIV) and Tuberculosis (TB) results in synergistic increases in their prevalence, morbidity, and mortality. The occurrence of both infections in the World is a great public health problem. The situation in Africa is not different. There is a looming threat of a pandemic emerging in Ghana as it has been in other African countries (GHS, 2006).

In an enlightened form, HIV is the virus that attacks and destroys the infection-fighting CD4 cells of the body's immune system. It is noteworthy that, loss of CD4 cells makes it very difficult for the immune system to fight infections. When the immune system is progressively damaged by HIV, the infected person becomes immune suppressed and is therefore exposed to other opportunistic infections, especially TB. The advanced form of HIV is called Acquired Immunodeficiency Syndrome (AIDS). HIV epidemic is one of the most destructive health crises of modern times destroying families and communities around the world. HIV is transmitted through the blood, semen, genital fluids, or breast milk of an infected person. Among the modes of transmission, unprotected sex or sharing drug injection equipment with an infected person, are the most common ways HIV spread (AIDS info facts sheet, 2012). Although tremendous researches have been conducted in the field of HIV/AIDS, there is no cure to it. However, there are steps one can take to delay the start of full blown AIDS and reduce it vulnerability. The



most promising advance made was the advent of potent combination of therapy in 1996. The antiretroviral therapy (ARV) drug's role is to prolong the life of the infected patient by slowing down wasting period by boosting the CD4 count in the immune system (Jackson, 2002).

On the other hand, TB is caused by the tubercle bacillus, Mycobacterium tuberculosis and spread through air (Friedland et al., 2007). TB attacks the lungs, but sometimes affects other parts of the respiratory system. TB that affects the lung is called the pulmonary TB; otherwise it is called extrapulmonary TB. There are two forms of TB; the latent TB and the TB disease. Latent TB infection is the inactive form; the TB germs in the body are sleeping and do not make a person sick. Others have strong immune system that quickly destroys the bacteria once they enter the body. A person with latent TB cannot spread it to unaffected persons. Without treatment, the latent TB infection can advance to TB disease (TB Facts, 2012). Generally, relatively small proportion of people infected with mycobacterium tuberculosis will develop TB disease, people who have much higher chances of developing the disease are those infected with HIV. TB cases are reported mostly among men than women, and affects adults in their productive ages. The burden of TB continues to increase due to poverty, population growth and HIV/AIDS (Tarimo, 2012). TB is the most common opportunistic infection complicating HIV infection especially in developing countries, and may occur at any stage in the course of immunodeficiency (Interagency Coalition on AIDS and Development, 2010).

HIV affects the immune system and increases the possibility of people acquiring new TB infections. It also promotes both the progression of latent



TB infection to active disease and a relapse of the disease in previously treated patients. TB is one of the leading causes of death for HIV infected people. HIV increases unfavorable drug reactions to treatment among TB patients. In addition, HIV infected people who have recovered from TB have an accelerated course of HIV disease and shortened survival compared with HIV infected people without a history of TB (WHO, 2004). People living with HIV (PLWH) are estimated to have between 12 to 20 times higher risk of developing TB disease compared to people living without HIV infection (Padmapriyadarsini *et al.*, 2011). It is also estimated that communities of higher HIV and TB rates have been rising severely even where effective TB control strategies are available. This underscored the need to improve TB control wherever HIV co-infection is common. The link between TB and HIV/AIDS may make people equate TB with HIV/AIDS. This may lead to increasing stigma and discrimination and delay TB patients in quest of health care and treatment (Refera, 2012).

HIV and TB pathogens potentiate each other accelerating the deterioration of immunological functions and resulting in premature death if untreated. Both TB and HIV have profound effects on the immune system as they are both capable of disarming the host's immune response through mechanisms that are not fully understood. Co-infection is the most powerful known risk factor for the progression of TB to active disease. HIV/TB co-infection significantly changes the original history of both diseases. This gives rise to different problems. One major problem associated with patients co-infected is intersecting both signs and symptoms between the diseases. The strange signs and symptoms of co-infection make the clinical diagnosis difficult in nearly all

cases. The fact that HIV/AIDS also makes the patient susceptible to other opportunistic infections with symptoms similar to TB are among the difficulties encountered in diagnosing TB in HIV patients. The other impact is that two different diseases with varying modes of diagnosis and treatment exist in a single patient. Both diseases involve the combination of different drugs. HIV/AIDS treatment is for life and a minimum of six month for TB (Pawlowski *et al.*, 2012).

Further, there has not been any efficient nationwide study on the prevalence of HIV/TB co-infection in Ghana. However, it is estimated that the influence of these diseases have been increasing such that in 1989 while about 14 percent of TB cases could be attributed to AIDS, by the year 2009 about 59 percent of the projected TB cases were attributed to the HIV/AIDS epidemic. Hospital studies have shown that the prevalence of HIV in TB patients is approximately 25-30 percent and that as many as 50 percent of patients with chronic cough could be HIV positive. Autopsies reports in Accra found that the proportion of TB deaths increased from 3.2 percent in 1987 to 1988 at the beginning of the HIV epidemic to 5.1 percent in 1997 to 1998. About 30 percent of PLWH at Korle-Bu Teaching hospital in 2007 were TB positive. TB accounts for 40-50 percent of HIV deaths in Ghana, while HIV is an important cause of medical deaths (GHS, 2007).

#### 1.2 Problem Statement

TB is communicable and airborne disease. It is the second leading cause of death from a single infectious agent, after the HIV. About 8.6 million people fell ill with TB in 2012, as well as 1.1 million cases among PLWH. In 2012,



1.3 million people died from TB, together with 320 000 among people who were HIV positive. Women who died from TB in 2012 approaches 410,000 including 160,000 among women who were HIV positive. Out of the overall TB deaths among HIV positive people, 50 percent were among women. TB is one of the top killers of women of reproductive age an estimated 530 000 children became ill with TB and 74 000 children who were HIV negative died of TB in 2012 (WHO Global TB Report, 2013). About half of all adults in Ghana carry a latent TB infection, which is suppressed by a healthy immune system. When the immune system is weakened by HIV, it can no longer control the TB infection and overt TB disease can develop. In the year 2000, approximately 11,300 new cases of TB were reported in Ghana. The TB Control Programme estimates the true figure to be more than 30,000. In 1989, about 14 percent of the TB cases could be attributed to AIDS. In 1997 study conducted in the Komfo Anokye Teaching Hospital Kumasi, Ghana found that HIV prevalence among TB patients was 23 percent HIV/AIDS epidemic (HIV/AIDS in Ghana, 2001).

Survival analysis is one of the appropriate techniques to demonstrate life time events and to identify the major prognostic factors. This method is appropriate because it assesses survival and the prognostic factors of each patient on treatment.

#### 1.3 Research Questions

This research will achieve its stated objectives if the following questions are duly answered.

i. What are the survival rates of the patients?

- ii. What predictor variables significantly influence the survival of patients?
- iii. What is the appropriate survival technique for this study?

#### 1.4 General Objective

The main objective is to evaluate the survival of HIV/AIDS, TB and coinfected patients and to identify the major prognostic factors that influence survival.

#### 1.5 The specific objectives

- i. To evaluate the mortality rate of HIV/AIDS, TB and co-infection patients
- ii. To examine the influence of prognostic factors on the survival of patients
- iii. To fit an appropriate survival model for the study

#### 1.6 Significance of the Study

The prevalence of HIV/AIDS, TB and co-infection among patients is high in African region. According to WHO report (2012), Ghana is among the forty-one countries with high HIV/TB burden. Ghana records 24 percent of her tested TB patients HIV positive. This high burden on the individual and the society lowers the productivity in the country.

This study will provide detailed knowledge about the prognostic factors and the most significant variables that have major impact on HIV, TB and HIV/TB co-infection patients and to identify the number of patients who died on treatment. It will be beneficial for policy makers and health workers to



institute effective measures to address the negative tendencies that these deadly diseases bring to the society. The study will be useful for further researches in the area of study.

#### **CHAPTER TWO**

#### LITERATURE REVIEW

#### 2.0 Introduction

The literature carefully chosen and reviewed in this chapter is relevant to the study. The literature is grouped into four thematic areas; Historical review of survival analysis, the review of: HIV/AIDS, TB and HIV/TB co-infection.

#### 2.1 Historical Review of Survival Analysis

The initial interest of survival analysis was death but now the scope has widened to include: time to the relapse of a disease, length of stay in a hospital, duration of a strike, money paid by health insurance, viral load measurements and time to finishing a thesis. Survival analysis is also used in the following studies: leukemia patients and time in remission, time to develop a heart disease for normal individuals, elderly population and time until death, and heart transplants and time until death products (Singh and Mukhopadhyay, 2011). The origin of survival analysis goes back to mortality tables from centuries ago. However, the new era of survival analysis emerged after the World War II. This was boosted as a result of the interest in the reliability of military equipment. This, resulted to the spread of private industry as customers became more demanding for safer and reliable products. As the usage of survival analysis grew in the field of clinical trials and medical researches, the parametric approach gave way for the nonparametric and semi parametric approaches. Survival analysis is suitable because medical intervention follow-up studies could start without all experimental units



enrolled at start of observation time and could end before all experimental units experience the event. Survival analysis is extremely imperative because some subjects will; withdraw, move too far away to follow, or die from some unrelated event. As such, censoring will enable researchers to analyse incomplete data (Smith and Smith, 2001). Censored data arises when we have information about individual survival time, but we do not know survival time exactly. Censoring is right censored, if it is known that the event of interest occurs sometime after the recorded follow-up period, left censoring occurs when the individual has experienced the event of interest prior to the start of the study, or interval censoring, where the only information is that the event occurs within some interval. Truncation schemes are left truncation, where only individuals who survive a sufficient time are included in the sample and right truncation, where only individuals who have experienced the event by a specified time are included in the sample (Klein and Moeschberger, 2003).

Furthermore, Chiang's (1960) expanded Kaplan and Meier's (1958) works aimed to deal with incomplete observations of survival data. Aalen's (1975) established mathematical theory of survival analysis based on the Martingale theory and Counting stochastic process. Aalen's work was not initially appreciated fully until the late 1980s. Fleming and Harrington (1991) and Andersen *et al.*, (1995) improved Aalen's work in 1975.

The development of statistical procedures and models for survival analysis improved marginally between 1970 and 1990 where survival analysis had established itself as an effective statistical method in biomedical research. Survival analysis became a dominant part of the standard biostatistics curriculum in medical schools, and now universally accepted for data analysis

in major medical fields. The Cox model was introduced by Cox, in 1972, for analysis of survival data with and without censoring, for identifying differences in survival due to treatment and prognostic factors in clinical trials. The Cox model is preferred over some statistical techniques including the logistic model because it ignores the survival time and censoring information. Given a Cox model and the coefficients, the baseline hazard function and the survival curves can be estimated (Singh and Mukhopadhyay, 2011).

#### 2.2 Some Reviews on HIV/AIDS

Renier's *et al.*, (2006) modelled the life table estimates of adult HIV/AIDS mortality in Addis Ababa Ethiopia. Between 54.7% and 62.4% of adult deaths in Addis Ababa (age 20 to 64) were attributed to AIDS. The study revealed that the absolute number of AIDS deaths in men is higher than women.

McMahon et al., (2011) modelled a prospective cohort study in poverty, hunger, education, and residential status impact survival in HIV. Their study examined data from 878 participants enrolled in NFHL studied from 1995 to 2005. The study took place at greater Boston Province in USA to investigate the effects of nutritional status of PLWH. The mortality rate was 23%, and median duration of follow up among the dead was 54.8 months. The study showed that age of the patients influence survival. The model deduced that patients with lower economic status are susceptible to death. Chi-square tests, Student's t-test, and Wilcoxon rank sum and Cox PH model were used.

Oduro and Aboagye-Sarfo (2011) researched into the modeling and control of HIV/AIDS propagation in the Ashanti Region of Ghana. Their study was carried out between 1982 and 2001 to assess the impact of the pandemic as well as the effectiveness of the existing control measures. Vector

Autoregressive time series analysis was employed to determine the discrete time linear autonomous models. The population dynamics of reported HIV/AIDS cases for females and males were found to be of second order, unstable, growing linearly in the mean but with a sinusoidal oscillation of period 4.2 years. The study revealed that condom use as a method of control has no significant impact on the disease.

Adams and Luguterah (2013) also studied the longitudinal analysis of change in CD4 cell counts of HIV patients on ARV in the Builsa District Hospital. Their study revealed that a patient's initial CD4 cell count significantly influence their present CD4 cell count. The duration of treatment was also significant. It was revealed that CD4 cell count increased in about 40 cells/mm³in every 6 months. This according to them suggests that there is strong positive association between CD4 count and duration of treatment. The study was estimated using the linear mixed effects model.

#### 2.3 Some Reviews on TB

Anyama et al., (2007) modelled the challenge of re-treatment pulmonary TB at two teaching and referral hospitals in Uganda. Their analysis discovered that the prevalence of re-treatment pulmonary TB at Mbarara based on medical records was 30% and 21.3% from exit interviews. The corresponding estimates at Mulago hospital were 12% and 43.9%. Compared to the 18 to 26 year age category, the Prevalence Odd Ratio for a seven year increase in age was 1.54, while female patients were 0.39 times less likely to report retreatment disease than males. A logistic regression was employed.

Pardeshi and Geeta (2009) worked on the survival analysis and risk factors for death in TB patients on DOTS. About 716 patients were registered at the TB unit. They recorded a survival rate at the end of the intensive phase of 96%, 93%, and 99% in the categories of I, II, III of DOTS respectively. There was no difference in the survival curves of male and female patients. Age groupings of 40 to 60 years and above 60 years were identified as significant risk factors. They employed Kaplan-Meier plot and log rank test to assess the survival pattern and Cox PH model.

Ponnuraja and Venkatesan (2010) studied survival models for exploring TB clinical trial data-an empirical comparison of PH model and AFT model. The data consists of 1236 TB patients admitted in randomised controlled clinical trial. They argued that the PH model displays significant lack of fit while the AFT model describes the data well.

Jakperik and Ozoje (2012) researched into the survival analysis of average recovery time of TB patients in Northern Region, Ghana. Their study was conducted on a retrospective moving cohort of sixty-one TB patients admitted into DOTS programme. Approximately 57.38% of the patients were males and 42.62% were females. New cases of TB were 88.52% and 11.48% relapse. Sixty-six (65.57) percent of the patients had Pulmonary TB, while 34.43% were diagnosed with extra Pulmonary TB. The study recorded 69% recovery and 31% treatment failures.

Jakperik and Kpakpo (2013) assessed the effects of prognostic factors in recovery of TB patients in the Upper West Region using Kaplan-Meier estimator and the logistic regression model. Out of the 400 patients they

studied over the period, 256 were males and 144 were females. It was also revealed that 62% of the respondents had pulmonary TB while 38% of the respondents had extra pulmonary TB. They recorded treatment success rate of 73.75% which was quite lower than the WHO target of 85%.

#### 2.4 Some Reviews on Co-infection

Murcia et al., (1996) evaluated the frequency of mycobacterium infection in an HIV positive population and its influence on medium term survival, along with clinical and epidemiological factors associated with co-infection. Several clinical specimens were studied for mycobacteria in a sample of 92 HIV positive patients at the San Juan de Dios teaching hospital in Bogota, Colombia, in 1996. Factors associated with infection were measured using a prevalence ratio at 95% confident interval. Logistic regression was used in the multivariable models. Eight percent (8%) of the patients had TB and 6% of them were found to be infected with atypical mycobacterium. Patients suffering from TB and stages III or IV HIV infection had a 16% survival rate.

Ngowi (2009) modelled HIV/AIDS and TB co-infection in rural Northern Tanzania. He sampled 440 patients, 102 health subjects for reference values, 105 newly diagnosed TB patients with unknown HIV status, and 233 PLWHA. The mean age for the PLWHA was (37.0  $\pm$  10.2). The overall HIV/AIDS and TB co-infection prevalence was 10.1%.

Mohammed *et al.*, (2011) conducted a case control study in Jimma and Mettu Karl hospitals where the two hospitals serve as referral and treatment centers for HIV and TB in south-west Ethiopia from January to March, 2009. Their study population consisted of 162 cases and 647 controls. PLWHA who

developed active pulmonary TB and controls were PLWHA without active TB. The result reveals after adjustment for potential confounders an initial weight less than 18.5kg a CD4 cell count less than 200 cells/mm³ a WHO clinical stage IV and not taking antiretroviral treatment were independently associated with the development of active TB in PLWHA. Multiple logistic regression was used.

Tarekegn (2011) conducted retrospective study in which a total of 632 patients, (316 in ART and pre-ART cohort) were followed for a median of 32.9 months in Pre-HAART and 35.4 months in HAART. The study was aimed to identify factors that increase the risk of TB in PLWHA. The result of the study indicated that WHO stage III or IV being bedridden and having hemoglobin level less than 10mg/dl were factors associated with increased risk of TB in PLWHA. Cox PH model was used.

Shaweno and Worku (2012) also employed a retrospective cohort study to compare the survival between HIV positive and HIV negative TB patients of 370 each, during an eight month DOTS period. They considered TB patients HIV status and follow up time until death was taken as an outcome. Cox PH regression model was used to determine the hazard ratio of death for each predictor. It had revealed that co-infected patients were less likely to survive.

Musenge et al., (2013) modelled the contribution of spatial analysis to understanding HIV/TB mortality in children using the structural equation modeling approach. They used multiple logit regression model with and without spatial household random effects. Structural equation models were also used in modeling the complex relationships between multiple exposures

and the outcome of HIV/TB child mortality. A protective effect was found in households with better socio-economic status and older children. Spatial models disclosed that the areas which experienced the greatest child HIV/TB mortality were those without any health facility.

Chu et al., (2013) studied the impact of TB on mortality among HIV patients on ART between 2000 and 2009 in Uganda using multiplicative Cox model. The percentages of death during follow-up were 10.47% and 6.38% for patients with and without TB, respectively. They discovered that HIV patients who had TB at the start of ART had an approximate 37% increased hazard of overall mortality relative to non TB patients.

Mor et al., (2013) worked on the TB incidence in HIV/AIDS patients in Israel from 1983 to 2010. They used a retrospective cohort study based on the National HIV and TB Registries. PLWHA who developed TB were compared to those who did not using the Cox model and Log rank test. The cumulative TB incidence among PLWHA in 2010 was 586 times higher than in HIV negative individuals. It was also revealed that, time for HIV patient to develop TB was shorter among males than in females.

#### 2.5 Conclusion

The chapter reviewed the literature that is relevant to the study. Review of the literature showed various techniques that researchers have employed in modelling HIV/AIDS, TB and HIV/TB co-infection survival data. However, among the various techniques reviewed the Cox Proportional Hazard model and the Accelerated Failure Model were used in this study.

#### CHAPTER THREE

#### RESEARCH METHODOLOGY

#### 3.0 Introduction

This chapter deals with the data and survival techniques that were used to achieve the stated objectives of this study. The chapter is sub-divided into eight sections including; source of data, study variables, descriptive statistics, comparison of survivorship functions, regression models, model selection criteria, model development and model diagnostics.

#### 3.1 Source of Data

The data for this study was obtained from St. Mathias Hospital in the Pru District of the Brong Ahafo Region of Ghana. This hospital serves as a referral center for different health centers in the District. The hospital has a unit for both ART and TB. The hospital started giving free ART services in 2008. Data was extracted from the patient folders, which have been adopted by the Ministry of Health, Ghana. The study considered all the patients on treatment with ages above five years. The study took place between 2008 and 2013 and the patient followed till the outcomes of either the event (treatment failure) or censored.

Pru District is one of the 27 districts in the Brong-Ahafo Region and it has Yeji as its administrative capital. It was originally curved out of the Atebubu-Amantin District in 2004 by an Act of Parliament through a Legislative Instrument and is the highest Administrative and Political authority within its

sphere of influence and jurisdiction. It is bordered to the north by East Gonja District in the Northern Region and to the south by Atebubu-Amantin and Nkoranza Districts. To the east, it shares boundaries with the Sene District and to the west with Kintampo South and Kintampo North Districts (http://pru.ghanadistricts.gov.gh/index.php).

#### 3.2 Study Variables

#### 3.2.1 Dependent Variable

The response or the experimental variable of this study is the survival time (months) from the day the patient begins ART and TB treatment till the day he/she dies or censor.

#### 3.2.2 Explanatory Variables

They predict changes on the dependent variables. In this study the following predictors are considered: Age in years, Weight (kg), Disclosure to Sexual Partner (no, yes), Marital status (Single, Married, Divorce, Widowed), Drug Regimen, Religion (Christian, Islam, and Traditionalist), Gender (Male, Female), Type of TB (Pulmonary, Extra-pulmonary) and WHO Clinical Stages (I, II, III, and IV).

#### 3.3 Descriptive Statistics

#### 3.3.1 Survivor Function S(t)

This measures the probability that a patient survives from time origin to sometime beyond t. It describes the proportion of the patients surviving to or beyond a given time. The actual survival time of a patient t, is regarded as the

value of a random variable T, which takes any non-negative value. The different values that T can take have probability distribution with its underlying probability density function f(t). The distribution function of T is given as:

$$F(t) = P(T < t) = \int_{0}^{t} f(u) du$$
 (3.1)

The survivor function S(t) is the probability that survival time is greater or equal t. That is:

$$S(t) = P(T \ge t) = 1 - F(t)$$
 (3.2)

As t ranges from 0 to  $\infty$  the survival function has the following properties, namely:

- \* it is non-increasing
- \* when t = 0, S(t) = 1. In other words, the probability of surviving past time 0 is sure.
- \*  $t \rightarrow \infty$ :  $S(t) \rightarrow 0$ . That is time goes to infinity, the survival curve approaches 0.

#### 3.3.2 Hazard Function h (t)

This is used to express the instantaneous failure rate. The probability that a patient dies at time t, conditioned that the patient survived. Thus, the probability that the random variable associated with a patient's survival time, T lies between t and  $t + \delta t$ , conditional on T being greater than or equal to t,  $P\{t \le T < t + \varphi t / T \ge t\}$ . This conditional probability is expressed as a probability per unit time by dividing the time interval by  $\varphi t$ , to give a rate. The

hazard function h(t) is the limiting value the quantity, as  $\varphi t$  tends to zero (Collet, 2003).

$$h(t) = \lim_{\varphi t \to 0} \left\{ \frac{P(t \le T < t + \varphi t / T \ge t)}{\varphi t} \right\}$$
 (3.3)

The hazard function in (3.3) can be expressed in terms of the probability density function and the survivor function as,

$$h(t) = \frac{f(t)}{S(t)} = -\frac{d}{dt} \{ \ln S(t) \}$$
 (3.4)

The cumulative hazard function H(t) is defined from (3.4) as,

$$H(t) = \int_0^t h(u) \, du = -\ln S(t) \tag{3.5}$$

#### 3.3.3 Survivorship Function Estimation

In this study, we estimated the survivorship function using the life table method also known as the actuarial method. The study employed the Gehan's method (1969) where the mid-points of the interval was used to estimate the hazard and the density functions and the upper limit used to estimate the survival function.

$$\hat{S}_{(t_i)} = \prod_{j=1}^{i-1} (1 - \hat{q}_j)$$
 (3.6)

$$\hat{f}_{(t_{mi})} = \frac{\hat{S}_{(t_i)} - \hat{S}_{(t_{i-1})}}{b_i} = \frac{\hat{S}_{(t_i)}\hat{q}_{(i)}}{b_i}$$
(3.7)

$$\hat{h}_{(t_{mi})} = \frac{d_i}{b_i \left( n_i - \frac{1}{2} d_i \right)} = \frac{2\hat{q}_i}{b_i - \hat{p}_i}$$
 (3.8)

in (3.6)-(3.8),  $t_{mi}$  is the mid-point of the  $i^{th}$  interval,  $d_i$  is the number of patients dying in the  $i^{th}$  interval,  $n_i$  is the number of patients exposed in the  $i^{th}$  interval,  $q_i = (d_i / n_i)$  is the conditional probability of dying in the  $i^{th}$  interval,  $\hat{p}_i = 1 - q_i$  is the conditional probability of dying in the  $i^{th}$  interval,  $b_i$  is the width of the  $i^{th}$  interval.

#### 3.4 Comparing Survivorship Functions

Having obtained the description of the general survival experience, the survival and hazard functions, we proceeded to compare the survivorship experience of the subgroups of qualitative variables in the data. These groups are defined by the values of covariates which are related to survival times. For easy interpretation, it is desirable that we graph each group concerned. The Kaplan-Meier estimator is appropriate for such graphs. A graph that shows the pattern of one survivorship function lying above another means a group well-defined by the upper curve have a longer survival time than the group defined by the lower curve. The plot will be significant if an appropriate statistical test is used (Hosmer and Lemeshow, 1999).

#### 3.4.1 Log Rank Test

The log rank test was used to compare the death rate between two distinct groups, conditional on the number at risk in the groups. This is a well-known and widely used test statistic. For k factor of groups, the log rank test hypothesis that;

 $H_0$ : All survival curves are the same,

 $H_1$ : Not all survival curves are the same.

Log rank test approximates a chi-square test which compares the observed number of failures to the expected number of failure under the hypothesis.

$$\chi^2 = \sum_{i=1}^k \frac{(O_i - E_i)^2}{E_i}$$
 (3.9)

In (3.9),  $O_i$  and  $E_i$  are the observed and expected number of death respectively. k-1 is the degree of freedom with k being the number of groups. A large chi-squared value will lead to the rejection of the null hypothesis in favour of the alternative.

#### 3.5 Regression Models for Survival Data

The data for each patient was collected to determine the relationship between the survival time and the covariates. This is to ensure the combination of the covariates that affects the hazard function. This will help obtain the estimates of the hazard function on the patients set of covariates.

#### 3.5.1 The Cox Proportional Hazards Regression Model

The Cox model is used to determine the effects the predictor variables have on the survival time. This model is usually written in terms of the hazard model formula. It defines the hazard at time t for a patient and a number of explanatory variables represented by X. The variable X represents a collection of predictor variables that is modeled to predict the patient's hazard. This is defined by;

$$h(t,X) = h_0(t) \exp\left(\sum_{i=1}^p \beta_i X_i\right)$$
(3.10)



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where,  $h_0$  (t) is the baseline hazard function,  $X_i$  is the explanatory or the predictor variable and  $\beta_i$  is the regression coefficients.

Cox (1972) proposed a semi-parametric model making it more robust to produce results that will closely approximate to a correct parametric model.

#### 3.5.2 Assumptions of the Cox proportional Hazard Model

i. The baseline hazard function  $h_0(t)$  depends on time t, but not covariates  $X_i$ 

ii. The hazard ratio  $\exp\left(\sum_{i=1}^{p} \beta_{i} X_{i}\right)$  depends on the covariates  $X_{i}$  not on time.

iii. The covariates  $X_i$  are time independent.

Assumption (ii) and (iii) can be expressed mathematically in terms of the hazard ratio, the hazard for one patient divided by the hazard for another patient. The two patients compared can be differentiated by their explanatory variables, that is, the X's.

$$\hat{HR} = \frac{\hat{h}(t, X^*)}{\hat{h}(t, X)} = \frac{h_0(t) \exp\left(\sum_{i=1}^{p} \hat{\beta}_i X_i^*\right)}{h_0(t) \exp\left(\sum_{i=1}^{p} \hat{\beta}_i X_i\right)}$$
(3.11)

$$\hat{HR} = \exp\left(\sum_{i=1}^{p} \hat{\beta}_i \left(X_i^* - X_i\right)\right)$$
 (3.12)

This apparently does not depend on time. This implies that the ratio of the hazard functions for two patients with different covariate values does not vary with time.

# 3.5.3 The Proportional Hazard Model Estimation The unknown coefficients of the Cox PH model

The unknown coefficients of the Cox PH model are estimated using the maximum likelihood. For the likelihood function to be applied to the survivorship function the Cox proportional data set are represented in three groups  $(t_i, \varphi_i, x_i)$ , i = 1, 2,...n.

 $t_i$  is the survival time for  $i^{th}$  person,  $\varphi_i$  is an indicator of censoring for the  $i^{th}$  patient given by 1 for censored and 0 for death,  $x_i$  a vector of covariates for patient  $i^{th}$   $(x_{il}, x_{i2}, ..., x_{ip})$ 

The full maximum likelihood is deduced as

$$L(B) = \prod_{i=1}^{n} h(t_i, x_i, \beta)^{\varphi i} S(t_i, x_i, \beta)$$
(3.13)

Where  $h(t_i, x_i, \beta) = h_0(t_i) \exp(\beta_i X_i)$  is individual hazard function for i,  $S(t_i, x_i, \beta) = S_0(t_i) \exp(\beta_i X_i)$  is individual survivorship function for i.

The full model becomes

$$L(B) = \prod_{i=1}^{n} \left( h_0(t_i) \exp(\beta_i X_i)^{\sigma} S_0(t_i) \exp(\beta_i X_i) \right)$$
(3.14)

We maximize (3.13) with respect to the unknown parameters of interest to obtain the full maximum likelihood.

#### 3.5.4 Partial Likelihood Function

Cox (1972) proposed the partial likelihood function that depends only on the parameter of interest. Thus, suppose that k of the survival time of the n patients are uncensored and distinct, and n-k are right censored. Consider



 $t_{(1)} < t_2 < \cdots < t_{(k)}, R(t_i)$  as the risk set at time.  $R(t_i)$  consist of all the persons whose survival time are at least  $t_i$ . For the particular failure at time  $t_i$ conditionally on the risk set  $R(t_i)$  the probability that the failure is on the individual as observed is

$$\frac{h(t,x)}{\sum_{j\in R(t_i)} h(t,x)} = \frac{h_0(t)\exp(\beta_i x_i)}{\sum_{j\in R(t_i)} h_0(t)\exp(\beta_i x_j)} = \frac{\exp(\beta_i x_i)}{\sum_{j\in R(t_i)} \exp(\beta_i x_j)}$$
(3.15)

This likelihood of the Cox PH model does not consider probabilities for all the patients on treatment.

$$L_{p}(\beta) = \prod_{i=1}^{m} \frac{\exp(\beta_{i} x_{i})}{\sum \exp(\beta_{i} x_{i})}$$
(3.16)

Where the product is over m distinct ordered failure times and  $x_i$  denotes the value of the covariate for the patients with ordered survival time  $t_i$ . The log partial likelihood function is

$$L_{p}(\beta) = \sum_{i=1}^{m} \left[ \beta_{i} x_{i} - \ln \left( \sum_{j \in R(t_{i})} \exp \left( \beta_{i} x_{j} \right) \right) \right]$$
(3.17)

The maximum partial likelihood (MPLE) can be obtained by differentiating (3.16) with respect to  $\beta_i$  setting the derivative to zero and solving the unknown parameters. This method is limited to only data sets that have no ties. Thus, there are no two or more variables with same survival time (Lee and Wang, 2003).



#### 3.5.4.1 Efron's Approximation

The problem of tied data set that makes the partial likelihood function estimation impossible is possible when the Efron's method is used. This method is considered to give a better result.

Let  $D_i$  be the set of subjects who are observed dead at time  $t_i$ .

$$L_{B}(\beta) = \prod_{j=1}^{m} \frac{\prod_{i \in D_{j}} \exp(\hat{X}_{i}\beta)}{\prod_{k=1}^{D_{j}} \left[ \sum_{l \in R_{j}} \exp(\hat{X}_{l}\beta) - \frac{k-1}{D_{j}} \sum_{l \in D_{j}} \exp(\hat{X}_{i}\beta) \right]}$$
(3.18)

The partial  $\log$  of (3.18) is given as;

$$L_{B}(\beta) = \sum_{i} \left\{ \sum_{i \in D_{j}} \beta \hat{X}_{i} - \sum_{l=0}^{k-1} \left\{ \log \left( \sum_{i:Y_{i} \geq t_{j}} \exp \left[ \beta \hat{X}_{i} - \frac{1}{k} \sum_{i \in D_{j}} \exp \left( \beta \hat{X}_{i} \right) \right] \right) \right\} \right\}$$
(3.19)

This is derived by differentiating (3.18) with respect to  $\beta$  component and equating it to zero.

#### 3.5.5 Accelerated Failure Time Model (AFT)

AFT models follow a known distribution. They are comprised of the Exponential model, Weibull, Lognormal, Log-logistics and Gamma model. The underlying assumption for this model is that the effect of the covariate is multiplicative with respect to the survival time. The AFT model is the natural logarithm of the survival time (*logt*). It is expressed as a linear function of the covariates.

$$\log t = \beta X_j + z_j \tag{3.20}$$

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Where,  $X_j$  is the vector of covariates,  $\beta$  is the vector of regression coefficient,  $z_j$  is the error term.

### 3.5.5.1 Exponential Distribution

This distribution is described as one parameter model because the hazard is constant over time. The risk of an event happening is flat over time. The hazard function is given as;  $h(t) = \lambda$ , with the cumulative hazard given as:

$$H(t) = \lambda t \tag{3.21}$$

For survival function we know that,  $H(t) = -\ln [S(t)]$ .

$$S(t) = e^{-H(t)} = e^{-\lambda t}$$
 (3.22)

For density function we multiply the hazard function by the survival function

$$f(t) = h(t)S(t) = \lambda e^{-\lambda t}$$
(3.23)

#### 3.5.5.2 Weibull Distribution

This model is flexible as compared to the exponential model because its hazard rates are not constant. It is a two-parameter model i.e.  $\lambda$  and p where,  $\lambda$  is the location parameter, p is the shape parameter. Thus, it informs whether the hazard is increasing, decreasing, or constant over time.

The hazard for the Weibull model is represented as;

$$h(t,X) = \lambda p(\lambda t)^{p-1}$$
 (3.24)

$$\lambda_i = e^{X_i \beta} \tag{3.25}$$

The shape parameter can be interpreted as:

If p < 1, then the hazard is monotonically decreasing with time.

If p > 1, then the hazard is monotonically increasing with time.

If p = 1, then the hazard is flat and we have the exponential model.

The survivor function is given as;

$$S(t) = e^{(-\lambda t)p} \tag{3.26}$$

The density function represented as;

$$f(t) = \lambda p(\lambda t)^{p-1} e^{(-\lambda t)p}$$
(3.27)

# 3.5.5.3 Log-Logistic Model

The hazard function for Log-Logistic is defined as;

$$h(t,X) = \frac{\frac{\lambda t}{\gamma} \left[ \frac{1}{\gamma} - 1 \right]}{\gamma \left( 1 + (\lambda t)^{\frac{1}{\gamma}} \right)}$$
(3.28)

$$\lambda_i = e^{-(X_i \beta)} \tag{3.29}$$

The log-logistic model have two parameters as the Weibull model,  $-\lambda$  been the location parameter and  $\gamma$  as the shape parameter. The hazard for Log-logistic is not monotonic. The shape parameter is defined as:

If  $\hat{\gamma}$  < 1then the conditional hazard first rises, then falls.

If  $\hat{\gamma} \ge 1$  then the hazard is declining

The survivor function for the log-logistic is

$$S(t) = \frac{1}{1 + (\lambda t)^{\gamma_r}} \tag{3.30}$$

The density function defined as;

$$f(t) = \frac{\frac{\lambda t}{\gamma} \left(\frac{1}{\gamma} - 1\right)}{\left\{\gamma \left(1 + (\lambda t)^{\frac{1}{\gamma}}\right)\right\}^{2}}$$
(3.31)

#### 3.5.5.4 Log-Normal Model

The survivor function for log-normal model is:

$$S(t) = 1 - \Phi\left\{\frac{\ln(t) - \mu}{\sigma}\right\} \tag{3.32}$$

Where  $\Phi$  is the standard Normal Cumulative distribution function and  $\sigma = \beta X$ . The density function denoted as:

$$f(t) = \frac{1}{\sigma t \sqrt{2\pi}} \exp\left\{-\frac{1}{2\sigma^2} (\ln(t) - u)^2\right\}$$
(3.33)

The hazard function for the Log-normal is given as:

$$h(t) = \frac{\frac{1}{\sigma t \sqrt{2\pi}} \exp\left\{-\frac{1}{2\sigma^2} (\ln(t) - u)^2\right\}}{1 - \Phi\left\{\frac{\ln(t) - \mu}{\sigma}\right\}}$$
(3.34)

The hazard rate for this model is similar to the log-logistic that is, where  $\hat{\gamma} < 1$  the hazard first rises and then falls.

# 3.5.5.5 The Gamma Model

The gamma distribution is a two parameter model with  $\lambda$  and k. The density function for the model is defined as:

$$f(t) = \frac{\lambda(\lambda t)^{k-1} e^{-\lambda t}}{\Gamma(k)}$$
 (3.35)

Survival function denoted as;

$$S(t) = 1 - I_k(\lambda t) \tag{3.36}$$

Where  $I_k(x)$  is the incomplete gamma function, and is represented as;



$$I_k(x) = \int_0^x \frac{\lambda^{k-1} e^{-x}}{\Gamma(k)} dx$$
 (3.37)

k is the shape parameter and  $1/\lambda$  the scale parameter.

The hazard function is obtained by taking the ratio of the density and the survivor function. i.e.  $\lambda(t) = \frac{f(t)}{S(t)}$ . The hazard for Gamma increases monotonically if, k > 1 from a value of 0 at the origin to a maximum of  $\lambda$ , is constant if k = 1 decreases monotonically if, k < 1, from 1 at the origin to an asymptotic value of  $\lambda$ .

#### 3.6 Model Selection Criteria

Akaike information criterion (AIC) and Bayesian information criterion (BIC) is used to compare the goodness-of-fit between the Cox model and the AFT model. These models are based on the log-likelihood l(b), the number of parameters in the distribution, p, and the total number of observations, n. Where  $\hat{b}$  denotes the MLE of all the parameters in the distribution. Models with smaller AIC or BIC values show a better fit. However, the BIC is preferred if the distribution have a sufficiently large sample size because it penalises models more severely than the AIC does.

$$BIC = l(\hat{b}) - \frac{p}{2}\log n \tag{3.38}$$

$$AIC = l(\hat{b}) - 2p \tag{3.39}$$

#### 3.7 Model Development

Since it is likely that the available covariates to be included in the model based on the clinical importance and statistical significance can be more than expected. It is necessary we decide on a reasonable method to select our covariates (Hosmer and Lemeshow, 1999). In this study, covariates that were insignificant were removed one at a time through the backward elimination system. The stepwise selection criterion was also used to select the covariates for the AFT model.

## 3.7.1 The Likelihood Ratio Test (LR)

The likelihood ratio (LR) is statistic used to test the significance of the interaction terms. This test statistic is computed by taking the difference between the log likelihood statistic of the reduced model which does not contain the interaction term and the log likelihood statistic of the full model containing the interaction. The decision rule to reject the null hypothesis is that:  $H_0$ : there is no interaction. Thus, if:  $LR = (2LL_R - (-2LL_F)) > \chi^2_{1,0.10} = 2.71$ . This test statistic approximates a chi-square distribution (Kleinbaum and Klein, 2005).

# 3.8 Assessing the Adequacy of the Models

Statistical inferences that lead to the identification of important risk or prognostic factors depend largely on the adequacy of the model selected. It is therefore very imperative to diagnose the Cox model to see whether it satisfies the proportionality assumption. Also, the AFT model is checked to ensure that it is well fitted. Thus, we employed the Cox-Snell residual, the martingale residual, Deviance residual and Schoenfeld residual.



#### 3.8.1. Cox-Snell Residual

This is instituted to estimate the overall fitness of the model. For the  $i^{th}$  patient the Cox-Snell residual is given as;

$$rz_i = \exp(\hat{\beta}x_i)\hat{H}_0(t_i) \tag{3.40}$$

 $\hat{H}_0(t_i)$ , is plot of the estimate of the baseline cumulative hazard function at time  $t_i$ , the observed survival time for each patient. If the fitted model is correct, the value  $rz_i$  will have a unit exponential distribution (Collet, 2003).

### 3.8.2 Martingale Residual

The martingale residuals have a mean of zero when the observations are uncensored. It takes values between  $-\infty$  and 1. The residuals sum to zero. In large samples the martingale residuals are uncorrelated and the expected values are zero. The properties of martingale residual are similar to the linear regression.

$$r_{Mi} = \varphi_i - rz_i \tag{3.41}$$

The quantity  $r_{Mi}$  is the difference between the observed numbers of deaths for the  $i^{th}$  patient in the interval (0, t). The interval (0, t) is the expected number of deaths.

 $r_{Mi}$  is an estimate of  $H_i(t_i)$  the cumulative hazard or the cumulative probability of death for the  $i^{th}$  patient (Collet, 2003).

#### 3.8.3 Deviance Residual

The martingale residuals are not symmetrically distributed about zero, even when the fitted model is correct making the result difficult to interpret. As a result the Therneau *et al.*, (1990) introduced the deviance residual which is more symmetrically distributed about the zero.

$$r_{Di} = sign(r_{Di}) \{ -2[r_{Di} + \varphi_i \log(\varphi_i - \varphi r_{Mi})] \}^{1/2}$$
 (3.42)

Where;  $sign(r_{Di})$  is a function that take values of +1 if the argument is positive and -1 if otherwise. This ensures that the variance residual have the same sign as the martingale residual. The original motivation for these residuals is that they are components of the deviance. The deviance is a statistic used to summarize the extent to which the fit of the model of current interest deviates from that of a model which is a perfect fit to the data. The statistic is given by;

$$D = -2\left(\log \hat{L}_c - \log \hat{L}_f\right) \tag{3.43}$$

Where,  $\hat{L}_c$  is the maximized partial likelihood under the current model and  $\hat{L}_f$  is the maximized partial likelihood of the full model. The model with a smaller deviance value is considered the best (Collet, 2003).

#### 3.8.4 Schoenfeld Residual

This method was proposed by Schoenfeld (1982) which differ from the previously mentioned. It is computed for the covariate of each patient. It is based on the first derivative of the log-likelihood function. Asymptotically, the Schoenfeld residuals have a mean of zero. These residuals will not be correlated with the covariates if the model is well fitted. A Schoenfeld residual



for the  $j^{th}$  covariate of the  $i^{th}$  patient with the observed survival time  $t_i$  is deduced as;

$$R_{ji} = \varphi_i \left\{ X_{ji} - \frac{\sum_{l \in R(t_i)} X_{jl} \exp(\hat{\beta} x_l)}{\sum_{l \in R(t_i)} \exp(\hat{\beta} x_l)} \right\}$$
(3.44)

Where,  $\hat{eta}$  is the maximum partial likelihood estimator of the Schoenfeld residuals are defined only at uncensored survival times, for censored observations they are set as missing. The sum of the Schoenfeld residuals for a covariate is zero.

# 3.9 Conclusion

The chapter dealt with the statistical techniques employed in this study. It presented the techniques in a clear, detailed and concise manner.





### DATA ANALYSIS AND DISCUSSION OF RESULTS

### 4.0 Introduction

This chapter analyses, interprets and discusses the results from the study. The chapter is grouped into Descriptive analyses, further analysis and discussion of results.

# 4.1 Descriptive Analyses

This segment explains the descriptive statistics of the data on the survival of HIV/AIDS, TB and the HIV/TB co-infected patients on treatment.

Table 4.1 revealed that a total of 590 patients at St. Mathias Hospital were on treatment. The patients were grouped into three, the HIV/AIDS, TB and HIV/TB co-infection. Of the 295 HIV/AIDS patients, 58 (19.7%) died; 219 were TB patients, 23 (10.5%) died and 76 were HIV/TB co-infection, 25 (32.9%) died during the period 2008 to 2013. It was also revealed that the coinfected patients experienced the worse survival followed by the HIV/AIDS patients. The TB patients were shown to have the best survival among all the three categories.

Table 4.1: Descriptive statistics for the three categories

Category	Total	Death	Censored	Percent Death
HIV/AIDS	295	58	237	19.7%
TB	219	23	196	10.5%
CO-INFECTED	76	25	51	32.9%
GRAND TOTAL	590	106	484	18.0%





Table 4.2 shows the summary statistics for continuous covariates of the patients. The maximum age of the HIV/AIDS patient was 75 years. The mean age of HIV/AIDS patient was 35 years. The median weight of the HIV/AIDS patients was 50 kg. The minimum and maximum weights were 8 kg and 90 kg respectively. The minimum survival time for HIV/AIDS patients was 1 month. The median and maximum survival times were 8 and 69 months respectively. The HIV/TB co-infected patients had a mean age of 37 years and a maximum age of 70 years. The minimum weight of the patients was 9 kg. The maximum weight was 93 kg. The mean and maximum survival time was 11 and 68 months respectively. Also, the minimum age of the TB patients was seven (7) years and the maximum age recorded was 102 years. The minimum weight of the TB patients was 15 kg which is little above the weight of HIV/AIDS and the co-infected patients. The maximum weight was 76 kg quite lower than the HIV/AIDS and the co-infection. The mean survival time was approximately five months. The maximum survival time for TB patients was 14 months.

Table 4.2: Descriptive statistics for continuous covariates for the patients

	Mean	Median	Std. Dev.	Min.	Max.
HIV/AIDS			<del></del>		
AGE	35.47	34.00	11.51	6.00	75.00
WEIGHT	50.98	50.00	11.22	8.00	90.00
TIME	17.07	8.00	19.30	1.00	69.00
<b>CO-INFECTION</b>					
AGE	37.09	35.500	14.96	6.00	70.00
WEIGHT	42.96	43.00	13.77	9.00	93.00
TIME	11.24	6.00	15.50	1.00	68.00
TB					
AGE	45.031	42.00	20.04	7.00	102.00
WEIGHT	45.12	45.00	11.14	15.00	76.00
TIME	4.92	5.00	2.36	1.00	14.00

Table 4.3 Hazard, density and survival estimates of HIV/AIDS patients on treatment

Mid-	Hazard	SE	Density	SE	Upper	Survival	SE
point			·		limit		
1	0.0249	0.0069	0.0243	0.0066	2	1.0000	0.0000
3	0.0098	0.0049	0.0092	0.0046	4	0.9514	0.0131
5	0.0117	0.0058	0.0107	0.0053	6	0.9329	0.0158
7	0.0096	0.0055	0.0086	0.0049	8	0.9114	0.0187
9	0.0104	0.0060	0.0092	0.0053	10	0.8942	0.0209
11	0.0074	0.0052	0.0064	0.0045	12	0.8757	0.0230
13	0.0000	0.0000	0.0000	0.0000	14	0.8629	0.0244
15	0.0000	00000	0.0000	0.0000	16	0.8629	0.0244
17	0.0046	0.0046	0.0039	0.0039	18	0.8629	0.0244
19	0.0000	0.0000	0.0000	0.0000	20	0.8550	0.0254
21	0.0000	0.0000	0.0000	0.0000	22	0.8550	0.0254
23	0.0055	0.0055	0.0047	0.0047	24	0.8550	0.0254
25	0.0181	0.0104	0.0150	0.0085	26	0.8456	0.0268
27	0.0065	0.0065	0.0053	0.0052	28	0.8156	0.0310
29	0.0000	0.0000	0.0000	0.0000	30	0.8051	0.0323
31	0.0216	0.0125	0.0170	0.0096	32	0.8051	0.0323
33	0.0157	0.0111	0.0120	0.0083	34	0.7710	0.0364
35	0.0168	0.0118	0.0123	0.0086	36	0.7471	0.0390
37	0.0275	0.0159	0.0194	0.0109	38	0.7224	0.0415
39	0.0204	0.0144	0.0137	0.0095	40	0.6837	0.0448
41	0.0110	0.0110	0.0071	0.0070	42	0.6564	0.0470
43	0.0240	0.0168	0.0149	0.0104	44	0.6421	0.0481
45	0.0132	0.0132	0.0080	0.0079	46	0.6122	0.0503
47	0.0147	0.0147	0.0086	0.0086	48	0.5963	0.0515
49	0.0000	0.0000	0.0000	0.0000	50	0.5791	0.0528
51	0.0164	0.0164	0.0093	0.0092	52	0.5791	0.0528
53	0.0179	0.0179	0.0098	0.0097	54	0.5604	0.0543
55	0.0204	0.0204	0.0108	0.0107	56	0.5407	0.0558
57	0.0000	0.0000	0.0000	0.0000	58	0.5191	0.0576
59	0.0690	0.049	0.0335	0.0224	60	0.5191	0.0576
61	0.0476	0.0476	0.0206	0.0198	62	0.4521	0.0669
63	0.0000	0.0000	0.0000	0.0000	64	0.4110	0.0723
65	0.0000	0.0000	0.0000	0.0000	66	0.4110	0.0723
67	0.0000	0.0000	0.0000	0.0000	68	0.4110	0.0723
69	0.0000	0.0000	0.0000	0.0000	70	0.4110	0.0723

Table 4.3 shows the hazard, density and survival estimates of HIV/AIDS patients. The life table estimates indicate that the  $59^{th}$  month after the patient was diagnosed of HIV/AIDS is the riskiest month as approximately 7% of the patients failed [ $\widehat{HR} = 0.068966$ ]. This is followed by the  $61^{st}$  month where approximately 5% failed. This is represented graphically in Figure 4.1.



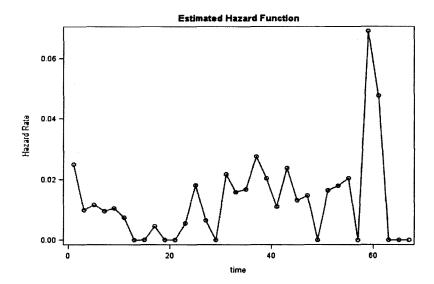


Figure 4.1: Hazard curve for HIV/AIDS

Table 4.4 shows the hazard, density and the survival estimates of the TB patients. The life table estimates revealed that about 3% of the risk occurred in the 1<sup>st</sup> month [ $\widehat{HR} = 0.029557$ ]. The 13<sup>th</sup> month was the riskiest month, as approximately half of the TB patients experienced the event [ $\widehat{HR} = 0.50000$ ], this is graphically represented in Figure 4.2.

Table 4.4: Hazard, density and survival estimates of TB patients on treatment

Mid- point	Hazard	SE	Density	SE	Upper limit	Survival	SE
1	0.0296	0.0085	0.0287	0.0081	2	1.0000	0.0000
3	0.0143	0.0064	0.0133	0.0059	4	0.9426	0.0161
5	0.0038	0.0038	0.00344	0.0034	6	0.9160	0.0196
7	0.0165	0.0117	0.0148	0.0103	8	0.9091	0.0206
9	0.0000	0.0000	0.0000	0.0000	10	0.8795	0.0286
11	0.1429	0.1414	0.1099	0.0953	12	0.8795	0.0286
13	0.5000	0.3062	0.2199	0.1102	14	0.6596	0.1916

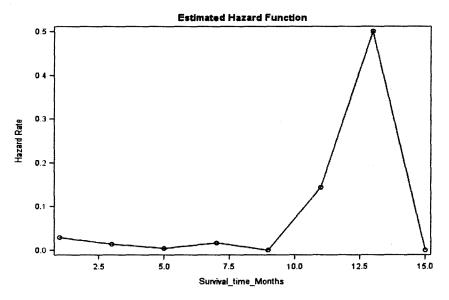


Figure 4.2 Hazard curve for TB Patients

About 7% of the co-infected patients  $[\widehat{HR} = 0.066667]$  failed in the first month of the treatment as shown in the Table 4.5 and Figure 4.3. The  $63^{rd}$  month was the riskiest month for the patients  $[\widehat{HR} = 0.333333]$ . Thus, approximately 33% of the patients failed.

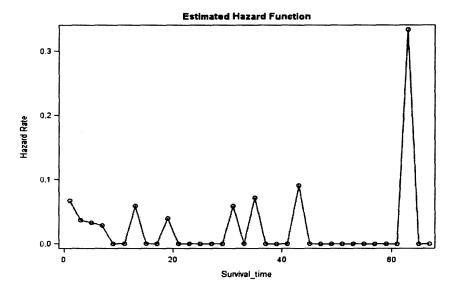


Figure 4.3: Hazard Curve for co-infection

Table 4.5: Hazard, density and survival estimates for co-infected patients on treatment

Mid-	Hazard	SE	Density	SE	Upper	Survival	SE
point					-limit		
1	0.0667	0.0222	0.0625	0.0195	2	1.0000	0.0000
3	0.0367	0.0183	0.0310	0.0150	4	0.8750	0.0390
5	0.0333	0.0192	0.0262	0.0147	6	0.8131	0.0469
7	0.0286	0.0201	0.0211	0.0146	8	0.7606	0.0528
9	0.0000	0.0000	0.0000	0.0000	10	0.7183	0.0577
11	0.0000	0.0000	0.0000	0.0000	12	0.7183	0.0577
13	0.0588	0.0415	0.0399	0.0268	14	0.7183	0.0577
15	0.0000	0.0000	0.0000	0.0000	16	0.6385	0.0739
17	0.0000	0.0000	0.0000	0.0000	18	0.6385	0.0739
19	0.0400	0.0399	0.0246	0.0238	20	0.6385	0.0739
21	0.0000	0.0000	0.0000	0.0000	22	0.5894	0.0829
23	0.0000	0.0000	0.0000	0.0000	24	0.5894	0.0829
25	0.0000	0.0000	0.0000	0.0000	26	0.5894	0.0829
27	0.0000	0.0000	0.0000	0.0000	28	0.5894	0.0829
29	0.0000	0.0000	0.0000	0.0000	30	0.5894	0.0829
31	0.0588	0.0587	0.0327	0.0312	32	0.5894	0.0829
33	0.0000	0.0000	0.0000	0.0000	34	0.5239	0.0962
35	0.0714	0.0712	0.0349	0.0331	36	0.5239	0.0962
37	0.0000	0.0000	0.0000	0.0000	38	0.4541	0.1057
39	0.0000	0.0000	0.0000	0.0000	40	0.4541	0.1057
41	0.0000	0.0000	0.0000	0.0000	42	0.4541	0.1057
43	0.0909	0.0905	0.0378	0.0356	44	0.4541	0.1057
45	0.0000	0.0000	0.0000	0.0000	46	0.3784	0.1120
47	0.0000	0.0000	0.0000	0.0000	48	0.3784	0.1120
49	0.0000	0.0000	0.0000	0.0000	50	0.3784	0.1120
51	0.0000	0.0000	0.0000	0.0000	52	0.3784	0.1120
53	0.0000	0.0000	0.0000	0.0000	54	0.3784	0.1120
55	0.0000	0.0000	0.0000	0.0000	56	0.3784	0.1120
57	0.0000	0.0000	0.0000	0.0000	58	0.3784	0.1120
59	0.0000	0.0000	0.0000	0.0000	60	0.3784	0.1120
61	0.0000	0.0000	0.0000	0.0000	62	0.3784	0.1120
63	0.3333	0.3143	0.0946	0.0725	64	0.3784	0.1120
65	0.0000	0.0000	0.0000	0.0000	66	0.1892	0.1450
67	0.0000	0.0000	0.0000	0.0000	68	0.1892	0.1450
69	0.0000	0.0000	0.0000	0.0000	70	0.1892	0.1450

In determining whether there is significant difference among different groups of the covariates, the log rank test of equality was employed as shown in Table 4.6. With the null hypothesis that: there is no significant difference between the survival curves of the patients.



Table 4.6: Test of equality using the log rank

Variable	df		$\chi^2$	<u> </u>		<i>p</i> -value	
		HIV	TB	Co-infected	HIV	TB	Co-infected
Gender	1	0.50	0.29	0.71	0.4811	0.5899	0.3991
Mstatus	3	1.92	0.93	4.36	0.5890	0.8185	0.2254
Religion	2	2.09	0.28	6.37	0.3520	0.6887	0.0414
WHO	3	39.17	*	1.62	0.0000	*	0.6555
Disclosure	1	0.66	*	0.39	0.8819	*	0.5313
Regimen	2	7.43	*	0.31	0.1901	*	0.8559
-		*	1.06	*	*	0.0040	*
TB Type	1	*	0.05	0.57	*	0.8287	0.4508

df: degrees of freedom

The log-rank test of equality shows a significant difference of survival among the groups; WHO Clinical Stage of HIV/AIDS patients, the Drug regimen of TB patients and Religion of the co-infected patients. However, covariates including Sex, Marital status, Religion of HIV and TB patients, Disclosure to sexual partner, Drug regimen of HIV/AIDS patients and TB type were not significantly different.

# 4.2 Further Analysis

# 4.2.1 The Cox Proportional Hazard Model for HIV/AIDS Patients

The proportional hazard model for HIV/AIDS patients in Table 4.7 showed that, the predictor variables Gender, WHO clinical stage and Weight are statistically significant at 10% significance level. The hazard estimate for Gender is given as  $[\widehat{HR}=0.4800, p\text{-value}=0.0370]$ . Thus, the result indicates that the rate of dying among female patients is approximately 50% lower than the male patients holding the other predictors constant. WHO clinical stages I, II and III of the patients showed an estimated hazard ratio and p-value as  $[\widehat{HR}=0.1640, p\text{-value}=0.0008], [\widehat{HR}=0.3080, p\text{-value}=0.0751]$  and  $[\widehat{HR}=0.3670, p\text{-value}=0.0179]$  respectively. This means that, the risk of death

<sup>\*:</sup> Means empty cell

for patients at WHO clinical stages I, II and III is about 16%, 30% and 36% respectively lower than the patients at WHO clinical stage IV assuming that all other predictors are constant. The estimated hazard ratio and p-value of weight is  $[\widehat{HR} = 0.9420, p < 0.0016]$  denoting that a unit change in the weight will reduce the risk of the patient by 0.9420 holding all other explanatory variables constant.

Table 4.7: The Cox proportional hazard regression model for HIV/AIDS

patients

Variables	Level	df	β	SE	$\chi^2$	<i>p</i> -value	Exp(β)
Gender	Female	1	-0.7349	0.3524	4.3492	0.0370	0.4800
Age		1	0.0066	0.0188	0.1238	0.7250	1.0070
	R	eligion	compared	with Trac	ditionalists		
Religion	Christian	1	-0.1587	0.6612	0.0576	0.8104	0.8530
_	Islam	1	-0.0387	0.7116	0.0030	0.9566	0.9620
	M.	[arital s	status comp	pared with	n Divorced		
Mstatus	Divorced	1	-0.6248	0.8084	0.5974	0.4396	0.5350
	Married	1	-0.3471	0.5455	0.4049	0.5246	0.7070
	Single	1	-0.0251	0.7373	0.0012	0.9728	0.9750
Weight	_	1	-0.0602	0.0191	9.9625	0.0016	0.9420
•		Regim	en compare	ed with C	BV/NVP		
AZT/3TC	/EFV	1	0.3207	0.4621	0.4818	0.4876	1.3780
		1	0.5308	0.4451	1.4224	0.2330	1.7000
AZT/3TC	/NVP						
	V	VHO cl	inical stage	e compare	ed with IV		
WHO	I	1	-1.8051	0.5359	11.3462	0.0008	0.1640
	II	1	-1.1780	0.6619	3.1674	0.0751	0.3080
	III	1	-1.0018	0.4233	5.6018	0.0179	0.3670
Disclosure	e No	1	-0.0750	0.5023	0.0223	0.8813	0.9280

df: degrees of freedom

In fitting the reduced model, the covariate that were insignificant at 10% significance level were removed one at a time from the model assessing at each stage the AIC values. Covariates including Disclosure, Regimen, Marital status and Religion were dropped from the full model with only Sex, Weight and WHO clinical stage retained as the only significant covariates as shown in appendix Table A1.

We assessed the importance of the insignificant covariates in the reduced Cox model for the HIV/AIDS patients to ensure that the spurious relationships are avoided. Thus, these covariates are added one at a time to the three significant covariates in the reduced model. The study revealed that none of those covariates were significant and therefore cannot be retained in the model. This implies that, the insignificant covariates are not as a result of confounding elements in the model as shown in appendix Table A2-A6.

Table 4.8: Reduced model for HIV/AIDS

Effects	Model	AIC
0	Sex weight WHO Disclosure Regimen MSTATUS	347.611
	Religion AGE	
1	Sex Weight WHO Disclosure Regimen Mstatus	343.739
	Religion	
2	Sex Weight WHO Disclosure Regimen Mstatus	338.900
3	Sex Weight WHO Disclosure Regimen	336.902
4	Sex Weight WHO Disclosure	335.076
5	Sex Weight WHO	332.965

In determining the interaction effects of the model, the possible interactions of the covariates were formed to see if their effects can increase or decrease the hazard rate of the patients. The study revealed that none of the covariates significantly interact at 10% significance level as shown in appendix Table A18. Therefore, the final model will be the model involving Sex, weight and WHO clinical stage.

For the model to be adequate for statistical inferences it is necessary to cross check the model with it assumptions. If the assumptions are duly met then the model is good enough for statistical predictions.

Table 4.9: Test of proportional hazard assumption

Time	rho	$\chi^2$	df	<i>p</i> -value
Gender	0.0971	0.34	1	0.5619
Age	-0.0061	0.00	1	0.9663
Religion	-0.0071	0.00	1	0.9667
Mstatus	0.0069	0.00	1	0.9607
Weight	-0.0944	0.24	1	0.6210
Regimen	-0.0870	0.24	1	0.6208
Who	0.1629	1.61	1	0.2039
Disclosure	-0.0918	0.31	1	0.5787
Global test		5.90	8	0.7502

df: degrees of freedom

In supporting the fact that the proportionality assumption was not violated, the Schoenfeld residual was performed as shown in Table 4.9. The correlation between the Schoenfeld residual for each of the covariate and the rank of the survival time was determined. The *p*-value and the global test were greater than the 5% significance level. Hence, we cannot reject the model that the proportionality assumption is violated. Thus, the Cox proportional hazard model is appropriate since all the covariates satisfied the proportionality assumption.

To validate this, we employed the graphical residuals to test if the assumption is duly met. The Scaled Schoenfeld residual in Figure 4.4 was used. This further suggests that there was no enough evidence that the Scaled Schoenfeld residual graphically violates the proportionality assumption. Undoubtedly, the plots support the proportionality assumption since the residuals are random and LOESS curves are smooth and horizontal with zero gradients.

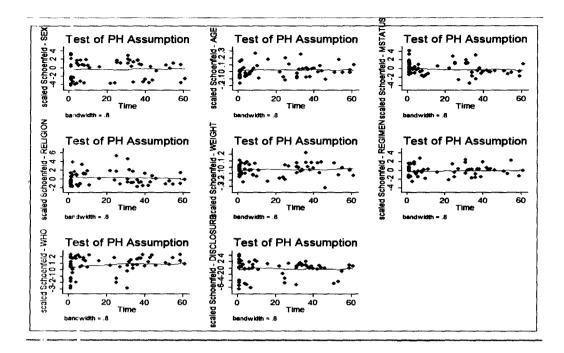
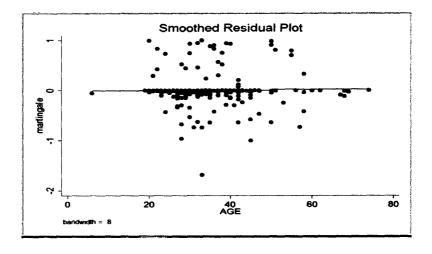


Figure 4.4 Scaled schoenfeld residual for HIV/AIDS covariates

The martingale residual plot is conducted to determine whether the assumption of the correct functional form of the model is satisfied and to establish whether the data support the hypothesis that the effect of the covariate is linear in the residual plot.



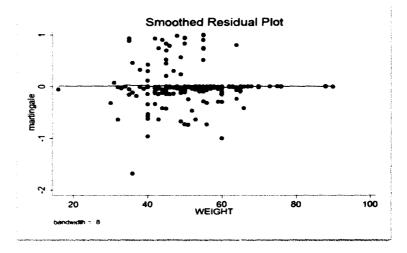


Figure 4.5: Martingale residual plot for continuous covariates of HIV/AIDS patients.

Figure 4.5 shows the plot of martingale residuals against the two continuous predictor variables Age and Weight. For each of the predictors the plot does not show trend and the resulting smoothed plots (LOESS) are approximately horizontal straight lines. Consequently, the martingale residual plot confirms that Age and Weight of the patients have a linear relationship with the survival time.

# 4.2.2 The Cox Proportional Hazard Model for Co-infection Patients

The proportional hazard model for the co-infected patients confirmed that the weight and Gender are significant as shown in Table 4.10. The estimated

hazard ratios are  $[\widehat{HR} = 0.9150, p - \text{value} = 0.0052]$  and  $[\widehat{HR} = 0.2980, p - \text{value} = 0.0415]$  respectively. This implies that an increase in the Weight of a patient will decrease the estimated hazard by 0.9150 assuming that all covariates are constant.

Table 4.10: Cox proportional hazard regression model for co-infection

patients		-1		<b></b>			-
Variables	Level	df	β	SE	$\chi^2$	<i>p</i> -value	Exp(β)
Gender	Female	1	-1.20975	0.59342	4.1559	0.0415	0.2980
Age		1	-0.0220	0.02166	0.0000	0.9872	1.0000
	R	eligion	compared	with Traditi	ionalists		
Religion	Christian	1	-0.0861	0.69072	0.0155	0.9008	0.9170
	Islam	1	0.7263	0.70620	1.0577	0.3037	2.0670
	N	<b>1arital</b>	status comp	pared with w	ridowed		
Mstatus	Divorced	1	0.5313	0.86656	0.3759	0.5398	1.7010
	Married	1	-0.7354	0.66948	1.2067	0.2720	0.4790
	Single	1	-4.3838	1.99875	4.8103	0.0283	0.0120
Weight		1	-0.0886	0.03169	7.8204	0.0052	0.9150
	Re	gimen	type comp	ared with Cl	3V/NVP		
AZT/3TC/	EFV	1	-0.3794	0.69116	0.3013	0.5831	0.6840
AZT/3TC/	NVP	1	-0.1414	0.57940	0.0596	0.8072	0.8680
	V	VHO c	linical stage	e compared	with IV		
WHO	I	1	-0.5896	0.61837	0.9092	0.3403	0.5550
	II	1	0.3599	0.78890	0.2082	0.6482	1.4330
	III	1	0.0369	0.67695	0.0030	0.9566	1.0380
Disclosure	No	1	-0.0109	0.58111	0.0004	0.9850	0.9890

df: degrees of freedom

Also, a female patient will have a decreased hazard of 0.2980 compared to the male patients assuming that all other covariates are constant. The Single patient is also significant  $[\widehat{HR} = 0.0120, p - \text{value} = 0.0283]$ . This implies that a Single patient have his/her estimated hazard decreased by approximately 98% compared to the widowed patient holding other factors constant.

The reduced model for the co-infection data was reached when the covariate that were insignificant at 10% significance level were removed one at a time from the model assessing at each stage the AIC values. Covariates including Disclosure, WHO clinical stage, Regimen, and Religion were dropped from



the full model with only Sex, Weight and Marital status maintained as the only significant covariates as shown in appendix Table A7.

Table 4.11: Reduced model for co-infection patients

Effects	Model	AIC
0	Sex Weight Mstatus Religion WHO Regimen	181.710
	Disclosure Age	
1	Sex Weight Mstatus Religion WHO Regimen	179.710
	Disclosure	
2	Sex Weight Mstatus Religion WHO Regimen	177.710
3	Sex Weight Mstatus Religion WHO	174.024
4	Sex Weight Mstatus Religion	169.654
5	Sex Weight Mstatus	169.323

The importance of the insignificant covariates in the reduced model for the coinfected patients was assessed to ensure that they do not confound the analysis. These insignificant covariates are added one at a time to the covariates in the reduced model. It is observed that none of those covariates were significant and therefore cannot be retained in the model. This implies that, the insignificant covariates are not as a result of confounding elements in the model as shown in appendix Tables A8-A12.

In assessing the interaction effects of the model, possible interactions of the covariates were formed to see if their effects can increase or decrease the hazard rate of the patients. The LR test was used to compare the log likelihood statistics for the interaction model and the no-interaction model. It is observed that none of the covariates significantly interact at 10% significance level to be included in the model as shown in appendix Table A19.

Table 4.12 shows the proportionality assumption for the co-infection. This provides enough evidence that the proportionality assumption is not violated

since the *p*-values are statistically insignificant. Hence, the Cox proportional hazard model is appropriate since all the covariates satisfied the proportionality assumption. This implies that the covariate does not correlate with the survival time.

Table 4.12: Test of proportional hazard assumption

Time	rho	$\chi^2$	df	<i>p</i> -value
Gender	0.1320	0.67	1	0.4141
Age	-0.0324	0.02	1	0.8901
Religion	0.1693	0.45	1	0.5037
Mstatus	0.1032	0.29	1	0.5884
Weight	0.2295	1.43	1	0.2316
Regimen	0.1009	0.34	1	0.5592
WHO	0.2242	1.18	1	0.2783
Disclosure	-0.1400	1.02	1	0.3128
Global test		12.00	8	0.2130

df: degrees of freedom

Figure 4.6 of the scaled Schoenfeld residual demonstrates that the proportionality assumption is duly satisfied. That is, since the residual plots are random and LOESS curves are smooth and horizontal with their slope being zero.

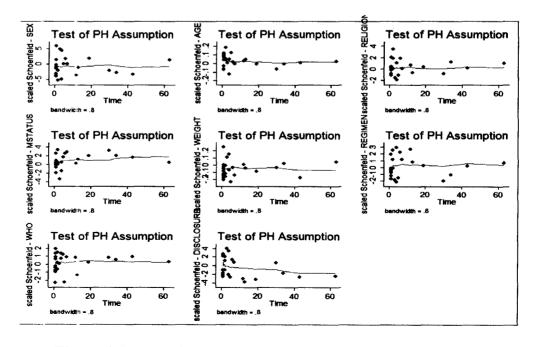


Figure 4.6 Scaled Schoenfeld residual for co-infected covariates



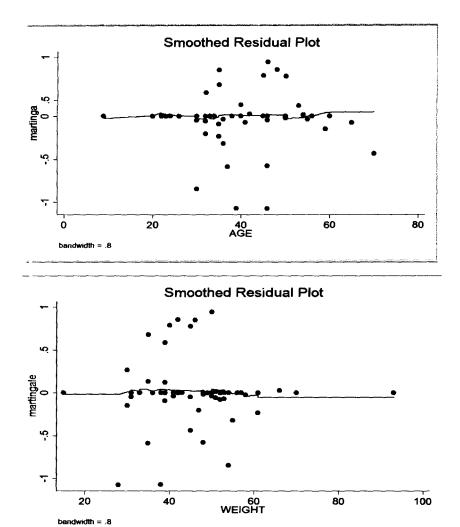


Figure 4.7: Martingale residual plot for continuous covariates for HIV/TB co-infection

We plotted the martingale residuals against two continuous predictor variables in the co-infection model: Age and Weight as shown in Figure 4.7. For each of the predictor variables, the plots showed a correct functional form. The result does not show any trend and the smoothed plots (LOESS) can be described as horizontal straight lines. The martingale residual plot confirms the model as having a linear relationship with the survival time. As such the model is appropriate.

# 4.2.3 The Cox Proportional Hazard Model for TB patients

Table 4.13 shows the fitted Cox proportional hazard model for TB patients. The weight and combine drug regimen shows a significant influence on the survival of the patients at 10% significant level. The estimated hazard ratio and p-value of the weight of patients is  $[\widehat{HR}=0.9600,\ p-\text{value}=0.0892]$ . This implies that a unit increase in the weight of a patient will eventually lower the risk of the patients by 0.9600 holding all other predictor variables constant. The estimated hazard ratios for drug regimen administered are  $[\widehat{HR}=0.024,\ p-\text{value}=0.0590]$  and  $[\widehat{HR}=0.3320,\ p-\text{value}=0.0773]$  for patients on the Drug regimens (HRZ) and (HRZE) respectively. This implies that as a patient uses this Drug regimen, the hazard of death decreases by 92% and 66% respectively compared to patients who are on (HRZES) drug regimen.

Table 4.13: Cox proportional hazard model for TB patients

1 able 4.13. Cox proportional hazard moder for 11b patients								
Variables	Level	df	β	SE	$\chi^2$	<i>p</i> -value	$Exp(\beta)$	
Gender	Female	1	0.30481	0.50217	0.3684	0.5439	1.3560	
Age		1	0.01345	0.01557	0.7463	0.3876	1.0140	
	Re	eligion	compared w	ith Traditio	nalists			
Religion	Christian	ĺ	-0.53001	0.64541	0.6744	0.4115	0.5890	
	Islam	1	-0.46550	0.70062	0.4414	0.5064	0.6280	
	Ma	arital s	status compai	red with Wi	dowed			
Mstatus	Divorced	1	-0.59622	1.1912	0.256	0.6167	0.5510	
	Married	1	-0.20446	0.7282	0.0788	0.7789	0.8150	
	Single	1	0.02914	1.1233	0.0007	0.9793	1.0300	
Weight		1	-0.04110	0.02418	2.8885	0.0892	0.9600	
TB type	Extra Pul	1	-0.36376	0.82393	0.1949	0.6589	0.6950	
	R	egime	n type compa	ared with H	RZES			
Regimen	HRZ	1	-2.83010	0.93189	9.2230	0.0024	0.0590	
	HRZE	1	-1.10377	0.62486	3.1203	0.0773	0.3320	

df: degrees of freedom

The reduced model for the TB data was reached when the covariate that were insignificant at 10% significance level were removed one at a time from the cox model assessing at each stage the AIC value. Among the covariates

removed included: Age, Sex, Religion, TB type and Marital status with only weight and Regimen retained as the only significant variables as shown in appendix Table A13.

Table 4.14: Reduced model for TB patients

Effects	Model	AIC
0	Weight Regimen Age Sex Religion	TB type Mstatus 215.910
1	Weight Regimen Age Sex Religion	TB type 210.345
2	Weight Regimen Age Sex Religion	208.538
3	Weight Regimen Age Sex	205.408
4	Weight Regimen Age	203.822
5	Weight Regimen	203.067

The insignificant covariates were assessed to ensure that the variables that will make the model bias are avoided in the TB data. Thus, these covariates were added one at a time to the reduced model. It is observed that, none of those covariates were significant and therefore cannot be retained in the model. This implies that, the insignificant covariates are not as a result of confounding elements in the model as shown in appendix Tables A14-A18.

In assessing the interaction effects of the model, possible interactions of the covariates were formed to see if their effects can increase or decrease the hazard rate of the patients. The LR compares the log likelihood statistics for the interaction model and the no-interaction model. The study revealed that none of the covariates significantly interact at 10% significant level as shown in Appendix A19. Therefore, interaction of the covariates will not be included in the model.

This shows that the proportionality assumption of the Cox PH model for TB patients is satisfied as shown in Table 4.15. The *p*-value is not significant at 5% significance level confirming that the proportionality assumption is satisfied

Table 4.15: Test of proportional hazard assumption

Time	rho	$\chi^2$	df	<i>p</i> -value
Gender	0.0638	0.10	1	0.7548
Age	-0.2237	1.21	1	0.2710
Mstatus	0.1415	0.56	1	0.4555
Religion	0.0460	0.04	1	0.8342
Regimen	0.1223	0.47	1	0.4923
Weight	-0.0300	0.03	1	0.8734
TB type	-0.0267	0.01	1	0.9089
Global test		2.36	7	0.9373

df: degrees of freedom

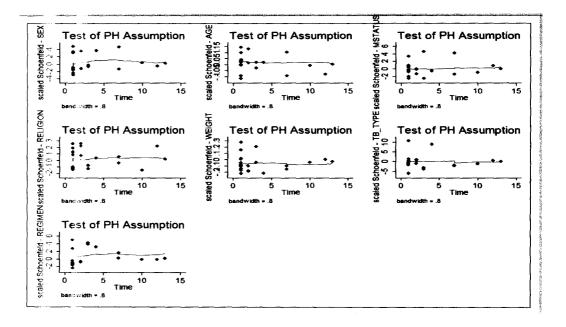
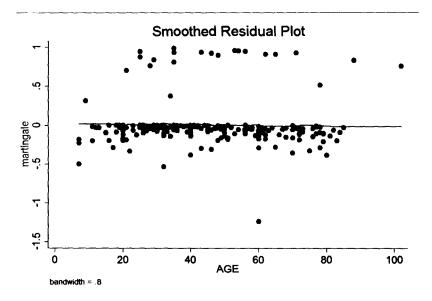


Figure 4.8 Scaled Schoenfeld residual for TB patients

Additionally, to affirm the adequacy of the proportionality assumption, the scaled Schoenfeld residual was duly examined in Figure 4.8. The residual plots are random and LOESS curves are smooth and horizontal with a slope zero. This suggests that the proportionality assumption is satisfied. Figure 4.9 is the plot of martingale residuals against the two continuous predictor variables age and weight in the TB Cox model. This is to aid us check for the linearity assumption. For each of the predictor variables age and weight, the plots display a correct functional form. It does not show trend and the resulting smoothed plots (LOESS) is described as horizontal straight lines. Thus, the

martingale residual plot confirms the model as having a linear relationship with the survival time. As such the model is adequate.



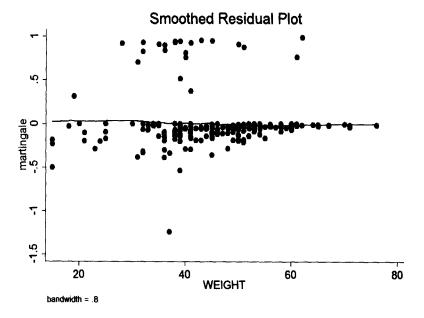


Figure 4.9: Martingale residual plot for continuous covariates of TB patients

# MND

# 4.3 Accelerated Failure Time Model (AFTM)

The HIV/AIDS and HIV/TB co-infection survival data shows that the AFT (Gamma) model is the best based on the AIC and BIC values. However, the TB data shows that the Cox model has the least BIC values as shown in Table 4.16.

Table 4.16: Model comparison

Criterion	Weibull	Exponent.	Gamma	Llogistic	Lnormal	Cox
HIV/AIDS	3					
AIC	273.362	274.138	246.477	278.993	280.761	347.611
BIC	328.510	325.839	305.072	334.141	335.908	370.901
Co-infection	o <b>n</b>					
AIC	135.472	134.551	129.336	136.916	136.007	181.710
BIC	172.764	169.512	168.958	174.207	173.299	198.774
TB						
AIC	198.394	196.750	200.839	197.777	195.669	215.970
BIC	242.451	237.419	248.286	241.829	239.727	228.461

The Gamma model shown in Table 4.17 revealed that Gender, Weight and WHO Clinical Stage are significant at 10% significance level. However, age, religion, regimen, marital status and disclosure of status to partner are insignificant.

The estimated survival time of a female patient will increase since the time ratio is greater than one  $[TR=e^{0.9427}=2.57]$ . Moreover, a unit increase in the weight of a patient will eventually increase the estimated survival time  $[TR=e^{0.0658}=1.09]$ . In other words, the survival time is 1.09 times longer for a unit increase in weight of a patient. Again, patients at WHO Clinical Stage I, II and III will have their predicted survival time accelerated by  $[TR=e^{2.3345}=10.32]$ ,  $[TR=e^{1.7459}=5.73]$  and  $[TR=e^{1.6392}=5.15]$  respectively.

Table 4.17: Gamma model for HIV/AIDS patients

Variables	Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -
, <b>441 14</b> 10 1410	20,07		Ρ				λ 	value
Intercept		1	0.7449	2.8028	-4.7484	6.2382	0.07	0.7904
Gender	Female	1	0.9427	0.4874	-0.0125	1.8979	3.74	0.0531
Age		1	-0.0089	0.0327	-0.0730	0.0552	0.37	0.7851
•		Rel	igion com	pared wit	h Tradition	nalist		
Religion		1	0.0319	0.8424	-1.6192	1.6829	1.34	0.2469
Christian								
	Islam	1	-0.1733	0.9317	-1.9993	1.6528	0.66	0.4166
		Maı	ital status	compared	l with Wid	lowed		
Mstatus	Divorced	1	0.5889	1.2211	-1.8044	2.9822	0.23	0.6296
	Married	1	-0.2117	0.8744	-1.9255	1.5021	0.06	0.8087
	Single	1	-0.5911	1.0712	-2.6906	1.5084	0.30	0.5811
Weight		1	0.0658	0.0349	-0.0027	0.1343	3.55	0.0597
		Re	gimen con	npared wi	th (CBV/N	IVP)		
AZT/3TC/	EFV	1	0.1568	0.5683	-1.2706	0.9570	0.08	0.7826
AZT/3TC/	NVP	1	0.7137	0.6589	-2.0050	0.5777	1.17	0.2787
		WH	O Clinical	l Stage co	mpared w	ith IV		
WHO	I	1	2.3345	0.8933	0.5836	4.0853	6.83	0.0090
	II	1	1.7459	0.9872	-0.1889	3.6807	3.13	0.0770
	III	1	1.6392	0.6535	0.3584	2.9200	6.29	0.0121
Disclosure	No	1	0.3030	0.6692	-1.0087	1.6147	0.21	0.6507
Scale		1	0.2209	0.0573	0.1328	0.3673		
Shape		_1_	7.2105	1.8461	3.5923	10.8288		

df: degrees of freedom

In fitting a reduced model for prediction, stepwise model selection was employed, with AIC criterion. The variables; Age, Regimen, Disclosure, Religion and Marital status were dropped with only Sex, Weight and WHO clinical stage as the significant covariates as shown in Table 4.18. Estimates of the reduced model are shown in Table 4.19.

Table 4.18 Stepwise model selection

Effects	Model	AIC
0	Sex Weight WHO Age Regimen Disclosure Mstatus	215.910
	Religion	
1	Sex Weight WHO Age Regimen Disclosure Mstatus	210.345
2	Sex Weight WHO Age Regimen Disclosure	208.538
3	Sex Weight WHO Age Regimen	205.408
4	Sex Weight WHO Age	203.822
5	Sex Weight WHO	203.067

Table 4.19 Estimate of reduced model

Variables	Level	df	β	SE	95%	6 C.I	$\chi^2$	<i>p</i> -value
Intercept		1	1.6664	1.2859	-0.8540	4.1868	1.68	0.1950
Gender	Female	1	0.3565	0.2966	-0.2247	0.9378	1.45	0.0393
Weight		1	0.0493	0.0297	-0.0089	0.1075	2.76	0.0967
_		WH	O Clinio	cal Stage	compared	with IV		
WHO	I	1 1	1.4427	0.6951	0.0803	2.8050	4.31	0.0379
	II	1 1	1.0510	1.8068	-2.4902	4.5923	0.34	0.5608
	III	1 (	0.9583	0.4620	0.0528	1.8639	4.30	0.0381
Scale		1 (	0.1881	0.0441	0.1188	0.2977		
Shape		1 9	9.4916	2.1905	5.1982	13.7849		

df: degrees of freedom

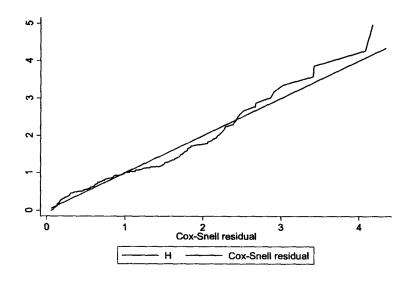


Figure 4.10 Gamma Cox-Snell residual for HIV/AIDS

The Cox-Snell residual graph as shown in Figure 4.10 compared to the rest of the AFT model in the appendix implies that the gamma graph is closer to the bisector than the rest of the models; as a result, the gamma model best fit the HIV/AIDS data in this study.

Table 4.20: Gamma model for co-infection patients

Table 4.	Table 4.20: Gamma model for co-infection patients									
Variable	es Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value		
Intercept	t	1	-2.7195	1.5647	-5.7862	0.3472	3.02	0.0822		
Gender	Female	1	0.7499	0.4391	-0.1108	1.6106	2.92	0.0877		
Age		1	0.0049	0.0012	0.0026	0.0072	17.56	<.0001		
			Re	eligion co	mpared wi	th Traditi	onalist			
	Christian	1	0.2927	0.5665	-0.8176	1.4031	0.27	0.6054		
	Islam	1	0.2875	0.6466	-0.9798	1.5547	0.20	0.6566		
			Ma	arital statu	is compare	d with W	idowed			
Mstatus		1	-0.2219	0.7405	-1.6734	1.2295	0.09	0.7644		
Divorce	1									
	Married	1	0.1907	0.5508	-0.8888	1.2703	0.12	0.7291		
	Single	1	2.7347	1.3184	0.1506	5.3188	4.30	0.0381		
Weight	_	1	0.0882	0.0240	0.0410	0.1353	13.44	0.0002		
		R	egimen con	mpared w	ith (CBV/I	NVP)				
AZT/3T	C/EFV	1	0.2338	0.6450	-1.0304	1.4980	0.13	0.7170		
AZT/3T	C/NVP	1	0.1261	0.5497	-0.9514	1.2035	0.05	0.8186		
		W	HO Clinica	al Stage co	ompared w	ith IV				
WHO	I	1	-0.4564	0.5366	-1.5081	0.5954	0.72	0.3951		
	II	1	-0.7908	0.6457	-2.0564	0.4749	1.50	0.2207		
	III	1	-0.3266	0.6711	-1.6420	0.9888	0.24	0.6265		
Disclosu	re No	1	-0.3633	0.5162	-1.3749	0.6484	0.50	0.4815		
Scale		1	0.6634	0.1070	0.4836	0.9100				
Shape		_ 1_	0.8089	0.1083	0.6222	1.0516				
	2.2 1									

df: degrees of freedom

The Gamma model shown in Table 4.20 indicated that Gender, Age and Weight were significant. However, Religion, Marital status, Drug regimen, WHO Clinical Stage and Disclosure were statistically insignificant.

The study reveals that female patients will survive longer with an estimated time ratio of;  $[TR = e^{0.7499} = 2.1167]$ . A unit increase in the age of a patient will accelerate his/her predicted survival time by  $[TR = e^{0.0049} = 1.0049]$ . Also, a unit increase in the weight of TB patient corresponds to an increase in the survival time by  $[TR = e^{0.0882} = 1.0922]$ .

In fitting a reduced model for prediction, stepwise model selection was employed, with AIC criterion. The variables; Age, Regimen, Disclosure, Religion and Marital status were dropped with only Sex, Weight and WHO clinical stage retained as the significant covariates as shown in Table 4.21.

Table 4.21: Stepwise model selection

Effects	Model	AIC
0	Mstatus Weight Age Sex WHO Religion Disclosure Regimen	129.336
1	Mstatus Weight Age Sex WHO Religion Disclosure	125.589
2	Mstatus Weight Age Sex WHO Religion	123.661
3	Mstatus Weight Age Sex WHO	119.922
4	Mstatus Weight Age Sex	117.874
5	Mstatus Weight Age	113.687

Table 4.22: Estimates of reduced model

Variables	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept	1	-0.3323	1.1709	-2.6273	1.9627	0.08	0.7766
Age	1	0.0033	0.0023	-0.0012	0.0078	2.08	0.0488
_		Marital :	status com	pared with	n Widowed		
Divorced	1	-0.9679	0.5165	-1.9802	0.0444	3.51	0.0609
Married	1	-0.0226	0.2773	-0.5662	0.5209	0.01	0.9349
Single	1	2.8596	0.7462	1.3971	4.3221	14.69	0.0001
Weight	1	0.0763	0.0283	0.0209	0.1317	7.29	0.0069
Scale	1	0.1401	0.0494	0.0702	0.2796		
Shape	1	10.6269	3.6709	3.4320	17.8218		

df: degrees of freedom

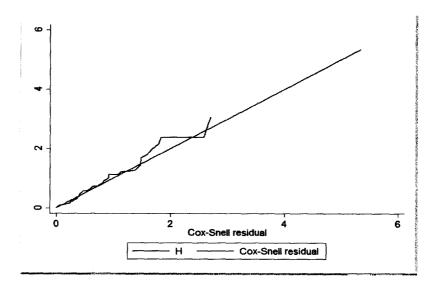


Figure 4.11 Cox-Snell residual plot for co-infection



The Cox-Snell residual graph as shown in Figure 4.11 implies that the Gamma model is the best fit for the co-infection data. This is because the model is closer to bisector than the rest of the models in the appendix.

#### 4.3 Discussion of Results

There were 590 patients on treatment from 2008 to 2013, of which two hundred and ninety-five (295) patients were diagnosed of HIV/AIDS; two hundred and nineteen (219) were TB patients and seventy-six (76) were HIV/TB co-infected. The study discovered that the percentage of death among the co-infected patients (32.9%) was higher than the deaths recorded from both HIV/AIDS and TB patients (30.2%). The survival estimates at the end of 70<sup>th</sup> month for HIV/AIDS and the co-infection was  $[\hat{S}_{(HIV/AIDS)} = 0.4110,$  $\hat{S}_{(Co-infection)} = 0.1892$ ] respectively. This suggests that the co-infected patients experienced the worse survival rate compared to the HIV/AIDS, evident in Table 4.2. This agrees with the Interagency Coalition on AIDS and Development in 2010 report, that 33% of all AIDS deaths worldwide can be attributed to TB. This could also be as a result of the difficulty in diagnosing the HIV patients of TB since HIV patients were more susceptible to contracting extra-pulmonary TB. It was also revealed that the youth are the most affected in HIV contraction with the average age of 35 years. The mean age of a TB patient was 45 years.

The hazard plot (Figure 4.1) for HIV/AIDS showed the  $59^{th}$  month of treatment as the riskiest month. About 50% of the risk of TB patients occurred in the  $13^{th}$  month [ $\widehat{HR} = 0.5000$ ]. This could be attributed to the Drug resistance strain resulting in hazard increase after the minimum recovery period of six month. The  $63^{rd}$  month of HIV/TB co-infection was the risky

month ( $\widehat{HR} = 0.3333$ ). This could be that the co-infected patients were not diagnosed of TB at the earlier state resulting in the higher risk at the later state of treatment.

In addition, the log rank test revealed that the WHO Clinical Stage is statistically significant for the HIV/AIDS patients. The Drug regimen for TB patients is also significant. There is a statistically significant difference among the religious denominations of HIV/TB co-infected patients. However, covariates including Gender, Marital status, Religion of both HIV and TB patients, Disclosure to sexual partner, Drug regimen for HIV/AIDS patients and TB type were statistically insignificant.

Further, Weight, Gender and WHO Clinical Stage of HIV/AIDS patients are clinically and statistically significant for the patient's survival in the Cox model. This suggests that, patients with higher Weight will have lower hazard hence an improvement in their survival rate. The result of the Gender showed female patients of recording a lower mortality rate than male. The patients at WHO clinical stage I, II and III have a better survival than the patient at stage IV. This could be as a result of the opportunistic infections such as the extrapulmonary TB and radiological bacterial pneumonia that the patients at WHO Clinical Stage IV usually surfer from.

The Weight, Gender and Marital status was also significant among the coinfected patients. The weight and the Drug regimen used were significant for the survival of the TB patients in the Cox model. Patients that used the combined drug regimen (HRZ) and (HRZE) have a better survival rate than the patients that used the drug regimen (HRZES). Thus, the patients on the (HRZ) and (HRZE) drug regimens are children and new cases of TB reported

respectively. Whereas, the patients on (HRZES) drug regimen are on retreatment due to relapse or the drug resistant strain hence the higher risk of survival.

The possible interactions of the covariates that were formed to see if their effects can increase or decrease the hazard rate of the patients revealed that none of the covariates significantly interact at 10% significant level.

Furthermore, comparing the Cox model with the Accelerated Failure Time model (AFT) showed that the AFT (Gamma) model was the best in the HIV/AIDS and HIV/TB co-infection survival data. However, the Cox model was adjudged the best model among the TB category based on the BIC values. In determining the prognostic effects among the HIV/AIDS patients using the Gamma model we observed that Gender, Weight and WHO Clinical Stage are statistically significant. The model deduced that the female patients had a better and longer predicted survival time. This is consistent with Owiti, (2013) where she argued that men naturally seek healthcare late and find it difficult to visit hospital regularly resulting in higher mortality rate among male. This suggests that a unit change in the Weight of patient will increase the predicted survival time. Rafera, (2012) asserts that the rate of dying among patients with higher weight in Ethiopia is proportionally lower compared to patients with lower Weight. Similarly, patients at WHO Clinical Stage I, II and III experienced an improved survival time. This could be as a result of the opportunistic infections at stage IV. The Gamma model for the co-infected patients showed that Gender, Age and Weight is statistically significant. Thus, an increase in the age of a patient will increase their predicted survival time.

The diagnostic evaluation of the Cox proportional hazard model proves that the proportionality assumption was satisfied. The global test revealed that there was no significant difference in the survival of the patients. Therefore, the proportionality assumption is satisfied. The scaled Schoenfeld residual plot was also performed to further justify the proportionality assumption. The test was adequate since the residual plots were random and LOESS curves smooth and horizontal with zero gradients. The martingale residual plot was also undertaken among the continuous covariates to check the correct functional form of the model. For each of the covariates, the plots do not show trend and the resulting smoothed plots (LOESS) were approximately horizontal straight lines. This confirms that the martingale residual plots have a linear relationship with the survival time. The Cox-Snell residual plot shows that the gamma model's graph is closer to the bisector than the rest of the models; as a result, the gamma model best fit the HIV/AIDS and HIV/TB co-infection survival data in the study.

#### 4.4 Conclusion

This chapter dealt with the analysis and discussion of results. It presented the major findings of the study in a clear, detailed, precise and concise manner.

#### **CHAPTER FIVE**

#### CONCLUSION AND RECOMMENDATIONS

#### 5.0 Introduction

This chapter covers the conclusion of the findings and some recommendations based on the study.

#### 5.1 Conclusion

In this study, the survival and the prognostic factors that affect HIV/AIDS, TB and HIV/TB co-infection between 2008 and 2013 were studied. The summary statistics and the survival estimate based on the life table showed that the HIV/TB co-infected patients experienced the worse survival rate.

Factors such as Weight, Gender and WHO clinical stage significantly determine the survival of the HIV/AIDS patients. We also observed that Gender, Age and Weight significantly determine the survival of the coinfected patients. While the Weight and the Drug regimen influenced the survival of the TB patients. Of the three categories the study deduced that Weight significantly determines the patient's survival.

The study showed the Cox proportional hazard model to be adequate for the TB survival data. However, the accelerated failure time model indicates that the Gamma model is well fitted for the HIV/AIDS and co-infection.

#### 5.2 Recommendations

Following the outcome of this study, the following recommendations were made:

- i. Government and stakeholders should support health institutions and physicians to initiate routine tests on opportunistic infections for HIV/AIDS patients to avoid deterioration in the health of the patients before the tests are conducted.
- ii. Health authorities and workers should be very cautious and pay much attention to patients who weighed lesser than the minimum weights of 8, 9 and 15 kilograms of HIV/AIDS, co-infection and TB respectively because it is observed in the study that this factor significantly affects the survival of the patients.
- iii. Researchers and authors who study in HIV/AIDS and TB coinfection should consider AFT (Gamma) model even if the proportionality assumption of the Cox model is satisfied.

#### REFERENCES

- Aalen, O. O. (1975): Statistical inference for a family of counting process.

  Ph.D. thesis, University of California, Berkeley.
- Adams, M., and Luguterah, A. (2013): Longitudinal Analysis of Change in CD4+ Cell Counts of HIV Patients on Antiretroviral Therapy (ART) in the Builsa District. *European Scientific Journal*, **33:** 1857 7881.
- AIDS info Fact sheet, (2012): available at <a href="http://aidsinfo.nih.gov">http://aidsinfo.nih.gov</a>
- Anyama, A., Sseguya, S., Okwera, A., El-Naggar, W, A., Mpagi, F., Owino, E. (2007): Modeled the challenge of re-treatment pulmonary TB at two teaching and referral hospitals in Uganda, *African Health Sciences* 7(3): 136-142.
- Andersen, P. K., O. Borgan, R. D. Gill, N. Keiding. (1993): Statistical Models based on Counting Process. Springer.
- Chiang C. L. (1960): A stochastic study of life tables and its applications: I.

  Probability distribution of the biometric functions. Biometrics,
  16:618-635.
- Chu R, Mills J. E., Beyene J., Pullenayegum E., Bakanda C., Nachega J.B.,
  Devereaux P. J., and Thabane L. (2013): Impact of TB on mortality
  among HIV patients on ART between 2000 and 2009 in Uganda.
  doi:10.1186/1742-6405-10-19
- Collett, D. (2003): Modeling Survival Data for Medical Research, Second Edition. Chapman and Hall.
- Cox, D. R. (1972): Regression models and life tables. J. R. Stat. Soc. Ser. B. 34:184-220.



- Cox, D. R. (1975): Partial likelihood. Biometrika 62:269-276.
- Cox, D. R. & D. Oakes. (1984): Analysis of Survival Data. Chapman and Hall, London.
- Fleming, T. R. & D. P. Harrington. (1991): Counting process and survival analysis. John Wiley and Sons. 429pp.
- Friedland G., Harries, A., and Coetzee, D. D. (2007): Implementation Issues in TB/HIV Program Collaboration and Integration. *The Journal of Infectious Diseases*, 196:S114–23. DOI: 10.1086/518664.
- Ghana Health Service, (2006): Technical Policy and Guidelines for TB/HIV Collaboration in Ghana.
- Ghana Health Service, (2007): Guidelines for the Clinical Management of TB and HIV co-infection in Ghana.
- Gehan, E. A. (1969): Estimating survival function from life table. *Journal of chronic diseases*. **21:** 629-644.
- HIV/AIDS in Ghana (2001): National AIDS/STI Control Programme. Disease Control Unit Ministry of Health, Ghana, 3<sup>rd</sup> edition.
- Hosmer, D.W., and Lemeshow, S. (1999): Applied Survival Analysis. John Wiley and Sons, Inc. New York.
- Interagency Coalition on AIDS and Development, (2010): Available at www.icad-cisd.com
- Jackson, H. (2002). AIDS Africa. Continent in crisis. Harare: SAFAIDS.
- Jakperik, D., and Aquaye, B. K. (2013): Assessing the Effects of Prognostic Factors in Recovery of Tuberculosis Patients in the Upper West

- Region. International Institute for Science, Technology and Education (IISTE), 11: 2224-5804.
- Jakperik, D., and Ozoje, M. (2012): Survival analysis of average recovery time of tuberculosis patients in Northern region, Ghana: *International Journal of Current Research*, **4:**123-125.
- Klein J. P., Moeschberger M. L., (2003): Survival analysis: techniques for censored and truncated data. Statistics for Biology and Health, Springer-Verlag New York, Inc.
- Kleinbaum, D. G. and Klein, M. (2005): Statistics for Biology and Health Survival Analysis. A Self-Learning Text: Springer.
- Lee, E. T., and Wang, J. W. (2003): Statistical Methods for Survival Data

  Analysis. 3rd Ed. New Jersey: John Wiley and Sons.
- Ma Z. and Krings A. W. (2008): Survival Analysis Approach to Reliability,

  Survivability and Prognostics and Health Management. Computer

  Science Department, University of Idaho Moscow, ID 83844, USA.
- McMahon, J., Wanke, C., Terrin N., Skinner, S., Knox, T. (2011): Poverty,

  Hunger, Education, and Residential Status Impact Survival in HIV.

  AIDS and Behavior, 15(7): 1503-1511.
- Mohammed, T. (2011): Risk factors of active tuberculosis in people living with HIV/AIDS in southwest Ethiopia: a case control study. *Ethiopian J Health Sci.*, **21**(2):131-9

- Mor, Z., Lidji, M., Cedar, N., Grotto, I., Chemtob, D. (2013) Tuberculosis incidence in HIV/AIDS patients in Israel from 1983-2010, PLOS, DOI: 10.1371/journal.pone.0079691.
- Musenge E., Vounatsou, P., Collinson, M., Tollman, S., and Kaln K. (2013):

  The Contribution of the Spatial Analysis to Understanding HIV/TB

  Mortality in Children a Structural Equation Modeling Approach.

  Global Health Action PMCID: PMC35566702 DOI: 103402/gha.v6i0.19266
- Ngowi B.J. (2009): HIV/AIDS and TB Co-infection in Rural Northern

  Tanzania: Epidemiology, clinical presentation and impact on CD4 cell

  counts. Centre for International Health University of Bergen,

  Norway. Unpublished PhD thesis.
- Oduro, F. T., Aboagye-Sarfo, F. (2011): Modeling and control of HIV/AIDS propagation: A case study in the Ashanti Region of Ghana. *Journal of Ghana Science*; No. 2
- Owiti, E. A. (2013): Cost Effectiveness and Survival analysis of HIV and AIDS Treatment in Kenya, University of Nairobi. PhD. Thesis.
- Padmapriyadarsini, C., Narendran, G., Swaminathan, S. (2011): Diagnosis & treatment of Tuberculosis in HIV co-infected patients, Indian Journal of Medical Research, 134(6): 850–865. DOI: 10.4103/0971-5916.92630



- Pardeshi and Geeta (2009): Survival Analysis and Risk Factors for Death in TB patients on Directly Observed Treatment-Short Course. *Indian Journal of Medical Sciences*; **No. 63** p180.
- Pawlowski A., Jansson M., Skold M., Rottenberg M.E., Kaillenius G.(2012)

  Tuberculosis and HIV Co-infection .journal.ppat. 8(2): e1002464
- Ponnuraja, C., Venkatesan, P. (2010): Survival models for exploring tuberculosis clinical trial data-an empirical comparison, *Indian Journal of Science and Technology*,7: 755-758.
- Pru District Assemly, available at <a href="http://pru.ghanadistricts.gov.gh/index.php">http://pru.ghanadistricts.gov.gh/index.php</a>, 2014
- Refera H., (2012): Survival and risk factors of HIV/TB Co-infected patients under antiretroviral therapy in Ambo hospital, Ethiopia. Addis Ababa University, School of Graduate Studies Department of Statistics. Unpublished M. Sc. Thesis.
- Reniers G., Araya T, Sanders E. J. (2006): Life table estimates of adult HIV/AIDS mortality in Addis Ababa. *Ethiop. J. Health Dev.* **20**(1): 3-9.
- Rich J. T., Neely J.G., Paniello R. C., Voelker C. C. J., DPhil, Nussenbaum B., and Wang E.W., (2010): A practical guide to understanding Kaplan-Meier curves. Otolaryngology-Head and Neck Surgery 143: 331-3366.
- Shaweno, D., and Worku A. (2012): Tuberculosis treatment survival of HIV positive patients on DOTS in Southern Ethiopia: A retrospective cohort study. *BioMed* 12; 5682.

- Singh R, Mukhopadhyay K., (2011): Survival analysis in clinical trials: Basics and must know areas. *Perspect Clin Res* 2:145-8.
- Smith, T. C., Smith, B. (2001): Survival analysis and the application of Cox's proportional hazards modeling using SAS. Proceedings of the 26th Annual SAS Users Group International Conference. Long Beach, CA Paper 244-26.
- Tarekegn, S. (2011): The effect of HAART on incidence of tuberculosis among HIV infected patients in Hawassa university referral hospital, South Ethiopia. Addis Ababa University. Unpublished Msc. Thesis.
- Tarimo, G. B. (2012): Delay in seeking care among tuberculosis patients attending tuberculosis clinics in Rungwe District, Tanzania.

  Unpublished Msc. Thesis.
- Tuberculosis Facts (2012): National Center for HIV/AIDS, Viral Hepatitis, STD, and TB Prevention. Division of TB Elimination.
- Toossi, Z, Xia, L, Wu M, and Salvekar, A. (1999): Transcriptional activation of HIV by Mycobacterium tuberculosis in human monocytes. *Clin Explmmunol*: 117: 324-30.
- Whalen, C., Horsburgh, C. R., Hom, D., Lahart C., Simberkoff, M., Ellner, J.J.
  (1995): Accelerated course of human immunodeficiency virus infection after tuberculosis. Am J Respir Crit Care Med; 151: 129-35

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- Whalen CC, Nsubaga, P., Okwerea, A. (2000): Impact of pulmonary tuberculosis on survival of HIV-infected adults: a prospective epidemiologic study in Uganda. 4:1219-28.
- World Health Organization Geneva, Switzerland (2004): Interim Policy on Collaborative TB/HIV Activities. Stop TB Department and Department of HIV/AIDS
- World Health Organization, (2013): Global tuberculosis report. Available at <a href="https://www.who.int/tb.">www.who.int/tb.</a>
- World Health Organization, (2013): Global update on HIV treatment: Results, impact and opportunities.

# 5

# APPENDIX I PUBLICATION

Nawumbeni, N. D., Luguterah, A., Adampah, T. (2014): Performance of Cox Proportional Hazard and Accelerated Failure Time Models in the Analysis of HIV/TB Co-infection Survival Data. *International Institute for Science Technology and Education*, **21**: 2224-5766.

# **APPENDIX II**

# TABLES AND FIGURES FOR COX MODEL FOR HIV/AIDS

# **PATIENTS**

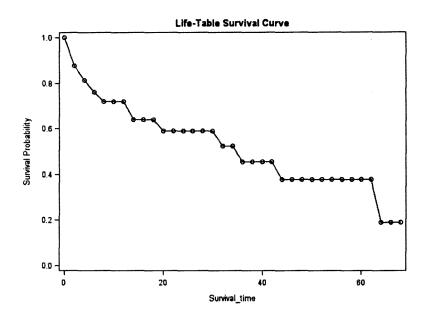


Figure A1: Survival curve for HIV/AIDS patients

Table A1: Reduced model									
Variable	;	df	β	SE	$\chi^2$	<i>p</i> -value	Exp(β)		
Gender	Female	1	-0.63551	0.32537	3.8150	0.0508	0.530		
Weight		1	-0.05591	0.01813	9.5124	0.0020	0.946		
		V	VHO clinical	stage comp	ared IV				
I		1	-1.69419	0.49097	11.9073	0.0006	0.184		
II		1	-1.04785	0.64047	2.6767	0.1018	0.351		
III		1	-0.84955	0.37430	5.1516	0.0232	0.428		



The insignificant covariates added one at a time to the reduced model.

Table A2: Age added

Testing Global Null Hypothesis: BETA=0							
Test	$\chi^2$	df	<i>p</i> -value				
Likelihood Ratio	33.2814	6	<.0001				
Score	36.0206	6	<.0001				
Wald	30.6384	6	<.0001				

Type III analysis of effects							
Effect	df	χ²	p-value				
Age	1	0.2689	0.6040				
Sex	1	3.6204	0.0571				
Weight	1	9.7245	0.0018				
WHO	3	12.6745	0.0054				

Table A3: Religion added

Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	33.3633	7	<.0001			
Score	36.7277	7	<.0001			
Wald	30.8192	_7_	<.0001			

Type III analysis of effects							
Effect	df	$\chi^2$	<i>p</i> -value				
Religion	2	0.3514	0.8389				
Sex	1	3.6583	0.0558				
Weight	1	9.2738	0.0023				
_WHO	_ 3	12.9346	0.0048				

Table A4: Marital status added

Testing Global Null Hypothesis: BETA=0						
Effect	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	33.0693	6	<.0001			
Score	35.9773	6	<.0001			
Wald	30.5789	6	<.0001			

Type III analysis of effects						
Effect	df	$\chi^2$	<i>p</i> -value			
Mstatus	3	1.4485	0.6942			
Sex	1	4.8066	0.0284			
Weight	1	9.2872	0.0023			
WHO	_3	14.3856	0.0024			



Table A5: Regimen added

Testing Global Null Hypothesis: BETA=0					
Test	$\chi^2$	df	<i>p</i> -value		
Likelihood Ratio	34.4491	8	<.0001		
Score	36.2176	8	<.0001		
Wald	30.0188	8	0.0002		

Type III analysis of effects							
Effect	df	$\chi^2$	<i>p</i> -value				
Regimen	2	1.2463	0.5362				
Sex	1	2.7820	0.0953				
Weight	1	6.0695	0.0138				
WHO	3	8.3091	0.0400				

Table A6: Disclosure added

Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	33.0693	6	<.0001			
Score	35.9773	6	<.0001			
Wald	30.5789	6	<.0001			

Type III analysis test							
df	$\chi^2$	<i>p</i> -value					
1	0.0515	0.8205					
1	3.8542	0.0496					
1	9.5380	0.0020					
3	13.4087	0.0038					
		df χ <sup>2</sup> 1 0.0515 1 3.8542 1 9.5380					



#### **APPENDIX III**

# TABLES AND FIGURES FOR COX MODEL FOR HIV/TB

# **CO-INFECTION PATIENTS**



Figure A2: Survival curve for Co-infection

Table A7: Reduced model co-infection								
Variables		df	β	SE	$\chi^2$	<i>p</i> -value	Exp(β)	
Gender	Female	1	-0.97975	0.46433	4.4522	0.0349	0.375	
	M	arital	status comp	pared with	widowed			
Divorced		1	0.44978	0.68316	0.4335	0.5103	1.568	
Married		1	-0.54453	0.53716	1.0276	0.3107	0.580	
Single		1	-4.01271	1.27652	9.8814	0.0017	0.018	
Weight		_ 1	-0.07735	0.02588	8.9330	0.0028	0.926	

The insignificant covariates added one at a time to the reduced model

Table A8: Age added

Testing Global Null Hypothesis: BETA=0						
Test $\chi^2$ df p-value						
Likelihood Ratio	17.6452	6	0.0072			
Score	12.8043	6	0.0463			
Wald	12.5717	6	0.0504			





Type III analysis test						
Effect	df	$\chi^2$	<i>p</i> -value			
Age	1	0.0956	0.7571			
Sex	1	4.4589	0.0347			
Mstatus	3	8.8360	0.0316			
Weight	1	8.9743	0.0027			

Table A9: Disclosure added

Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	17.5662	6	0.0001			
Score	12.8713	6	0.0001			
Wald	12.5245	6	0.0012			

Type III analysis test							
Effect	df	$\chi^2$	<i>p</i> -value				
Disclosure	1	0.0151	0.9021				
Sex	1	4.4314	0.0353				
Mstatus	3	11.0102	0.0117				
Weight	1	8.2187	0.0041				

Table A10: Religion added

Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	21.2193	7	0.0035			
Score	17.1290	7	0.0166			
Wald	14.3141	7	0.0459			

Type III and	alysis test		
Effect	df	$\chi^2$	<i>p</i> -value
Religion	2	4.0960	0.1290
Sex	1	4.2498	0.0393
Mstatus	3	10.2395	0.0166
Weight	1	8.2340	0.0041

Table A11: WHO added

Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	20.4200	8	0.0089			
Score	13.8725	8	0.0851			
Wald	12.7230	8	0.1217			

Type III analysis test						
Effect	df	$\chi^2$	<i>p</i> -value			
WHO	3	2.5724	0.4623			
Sex	1	5.9679	0.0146			
Mstatus	3	10.6039	0.0141			
Weight	1	9.5509	0.0020			

# Table A12: Regimen added

Tosting Clohal Null Hynothesis: BETA=0							
Likelihood Ratio	18.1331	7	0.0114				
Score	12.6946	7	0.0799				
Wald	12.6400	7	0.0814				

Type III ar	nalysis te	st	
Effect	df	$\chi^2$	<i>p</i> -value
Regimen	2	0.5936	0.7432
Sex	1	4.5735	0.0325
Mstatus	3	11.3586	0.0099
Weight	1	8.9828	0.0027



#### APPENDIX IV TABLES AND FIGURES FOR COX

#### MODEL FOR TB PATIENTS

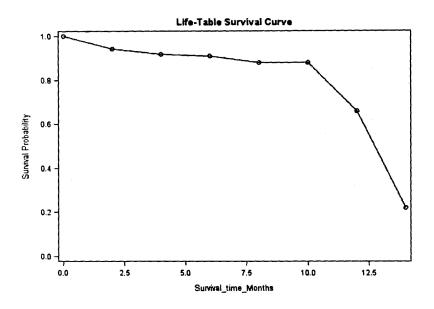


Figure A3: Survival curve for TB patients

Table A13: Reduced model for TB

Variable		df	β	SE	$\chi^2$	<i>p</i> -value	Exp(β)
Weight		1	-0.03462	0.01864	3.4493	0.0633	0.966
Regimen type compared with HRZES							
Regimen	HRZ	1	-2.54039	0.89020	8.1437	0.0043	0.079
Regimen	HRZE	1	-0.69685	0.52833	1.7397	0.1872	0.498

The insignificant covariates added one at a time to the reduced model

Table A14: Marital status added

Testing Global Null Hypothesis: BETA=0					
Test	$\chi^2$	df	<i>p</i> -value		
Likelihood Ratio	15.8786	6	0.0144		
Score	15.1595	6	0.0191		
Wald	12.8979	6	0.0447		

Type III ar	nalysis test		
Effect	df	$\chi^2$	<i>p</i> -value
Mstatus	3	1.2528	0.7404
Weight	1	2.9185	0.0876
Regimen	2	8.3768	0.0152



Testing Global Null Hypothesis: BETA=0						
Test	$\chi^2$	df	<i>p</i> -value			
Likelihood Ratio	14.8956	4	0.0049			
Score	14.6059	4	0.0056			
Wald	12.5293	4	0.0138			

Type III analysis tes	t		
Effect	df	$\chi^2$	<i>p</i> -value
TB type	1	0.1398	0.7085
Weight	1	3.5616	0.0591
REGIMEN	2	8.3916	0.0151

Table A16: Religion added

Testing Global Null	Hypothesis:	BETA	=0
Test	$\chi^2$	df	p-value
Likelihood Ratio	16.1692	5	0.0064
Score	15.7532	5	0.0076
Wald	13.4479	5	0.0195

lysis test		
df	χ²	<i>p</i> -value
2	1.6126	0.4465
2		0.0623 0.0131
		$df$ $\chi^2$

Table A17: Gender added

Testing Global Null Hypothesis: BETA=0						
Test Likelihood Ratio	$\chi^2$	df	p-value			
Score Score	14.9824	4	0.0047			
Wald	14.6292	4	0.0055			
	12.4685	4	0.0142			

Type III Effect		test	
Gender Weight Regimen	<b>df</b> 1 1 2	<b>0.2408</b> 3.1944 8.1994	<i>p</i> -value 0.6236 0.0739 0.0166

Table A18: Age added

Testing Global Null Hypothesis: BETA=0									
Test $\chi^2$ df p-value									
Likelihood Ratio	15.9917	4	0.0030						
Score	15.0804	4	0.0045						
Wald	12.8190	4	0.0122						

_Type III a	nalysis test		
Effect	df	$\chi^2$	<i>p</i> -value
Age	1	1.2746	0.2589
Weight	1	4.1040	0.0428
Regimen	2	8.4816	0.0144

**Table A19: Interaction effects** 

Interaction	$-2LL_R$	$-2LL_F$	$-2LL_R-(-2LL_F)$	DECISION
HIV/AIDS			,	
Sex*Weight	525.149	524.990	0.159	Fail to reject
Sex*WHO	500.145	497.600	2.545	Fail to reject
Weight*WHO	495.987	493.747	2.240	Fail to reject
CO-INFECTION				_
Sex*Mstatus	170.252	167.603	2.649	Fail to reject
Sex*WEIGHT	175.572	173.318	2.254	Fail to reject
Weight*Mstatus	163.922	161.448	2.474	Fail to reject
TB				v
Weight*Regimen	197.067	195.871	1.196	Fail to reject



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# APPENDIX V TABLES AND FIGURES FOR AFT MODEL

# FOR HIV/AIDS PATIENTS

Table A20: Weibull model

Variables Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value
Intercept	1	-1.5506	2.5841	-6.6155	3.5142	0.36	0.5485
Gender Female	1	0.8039	0.4767	-0.1305	1.7383	2.84	0.0918
Age	1	-0.0024	0.0287	-0.0585	0.0538	0.01	0.9337
	Re	ligion com	pared wit	h Traditio	nalist		
Christian	1	2.0442	0.9364	0.2089	3.8795	4.77	0.0290
Islam	1	1.7962	1.0033	-0.1703	3.7626	3.20	0.0734
	Maı	rital status	compared	d with Wic	lowed		
Mstatus Divorced	1	0.4206	1.0482	-1.6337	2.4750	0.16	0.6882
Married	1	0.8233	0.7924	-0.7297	2.3764	1.08	0.2988
Single	1	-0.3782	0.9909	-2.3203	1.5639	0.15	0.7027
Weight	1	0.0433	0.0272	-0.0100	0.0966	2.53	0.1115
	Re	gimen con	npared wi	th (CBV/N	VVP)		
AZT/3TC/EFV	1	-0.5295	0.6592	-1.8215	0.7625	0.65	0.4218
AZT/3TC/NVP	1	-1.2319	0.6632	-2.5316	0.0679	3.45	0.0632
	WH	IO Clinica	l Stage co	mpared w	ith IV		
WHO I	1	0.4217	0.7180	-0.9855	1.8290	0.35	0.5569
II	1	0.5910	0.8599	-1.0943	2.2763	0.47	0.4919
III	1	1.1856	0.5591	0.0897	2.2814	4.50	0.0340
Disclosure No	1	0.1199	0.7150	-1.2815	1.5213	0.03	0.8668
Scale	1	1.2362	0.1655	0.9509	1.6072		
Shape	1	0.8089	0.1083	0.6222	1.0516		

Table A21: Exponential model

Variables Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value
Intercept	1	-1.0675	2.1343	-5.2506	3.1157	0.25	0.6170
Gender Female	1	0.7438	0.3887	-0.0181	1.5057	3.66	0.0557
Age	1	-0.0030	0.0238	-0.0496	0.0436	0.02	0.8988
-	R	eligion co	mpared w	rith Traditi	onalist		
Christian	1	1.8633	0.7646	0.3648	3.3619	5.94	0.0148
Islam	1	1.5467	0.8107	-0.0423	3.1357	3.64	0.0564
	M	arital statu	s compar	ed with W	idowed		
Divorced	1	0.4488	0.8613	-1.2393	2.1370	0.27	0.6023
Married	1	0.8142	0.6499	-0.4597	2.0881	1.57	0.2103
Single	1	-0.3383	0.8039	-1.9139	1.2374	0.18	0.6739
Weight	1	0.0361	0.0220	-0.0070	0.0792	2.70	0.1006
	R	egimen co	mpared v	vith (CBV	/NVP)		
AZT/3TC/EFV	1	-0.4863	0.5418	-1.5483	0.5756	0.81	0.3694
AZT/3TC/NVP	1	-1.0408	0.5339	-2.0873	0.0057	3.80	0.0513
	W	HO Clinic	al Stage o	compared v	with IV		
WHO I	1	0.4218	0.5989	-0.7520	1.5956	0.50	0.4813
II	1	0.4747	0.6946	-0.8866	1.8360	0.47	0.4943
III	1	1.0732	0.4577	0.1761	1.9703	5.50	0.0190
Disclosure No	1	0.1772	0.5955	-0.9899	1.3443	0.09	0.7661
Scale	1	1.0000	0.0000	1.0000	1.0000		
Weibull scale	1	1.0000	0.0000	1.0000	1.0000		<u> </u>

Table A22: Log-normal model

Parameter	df	β	SE	95%	CI	$\chi^2$	<i>p</i> -value
Level	uı	Р	SE	73 /0	C.I	X	p-value
		0.0600	0.5506	7.4600	50445	0.00	0.0000
Intercept	1	-0.0623	2.7586	-5.4692	5.3445	0.00	0.9820
Gender Female	1	0.8077	0.5670	-0.3035	1.9190	2.03	0.1543
Age	1	-0.0112	0.0302	-0.0704	0.0480	0.14	0.7110
	F	Religion co	mpared w	ith Traditic	nalist		
Christian	1	1.5201	1.0090	-0.4575	3.4978	2.27	0.1319
Islam	1	1.3939	1.0542	-0.6722	3.4601	1.75	0.1861
	N	Iarital stat	us compar	ed with Wi	dowed		
Divorced	1	0.0269	1.1970	-2.3191	2.3730	0.00	0.9821
Married	1	0.5114	0.8915	-1.2358	2.2586	0.33	0.5662
Single	1	-0.6922	1.1622	-2.9701	1.5856	0.35	0.5514
Weight	1	0.0428	0.0322	-0.0204	0.1060	1.76	0.1844
	I	Regimen c	ompared v	vith (CBV/	NVP)		
AZT/3TC/EFV	1	-0.8269	0.7625	-2.3214	0.6675	1.18	0.2781
AZT/3TC/NVP	1	-1.2465	0.7132	-2.6444	0.1513	3.05	0.0805
	V	/HO Clini	cal Stage o	compared w	ith IV		
WHO I	1	0.3636	0.7916	-1.1880	1.9152	0.21	0.6460
II	1	0.6811	0.9231	-1.1281	2.4902	0.54	0.4606
III	1	0.9965	0.6813	-0.3388	2.3319	2.14	0.1436
Disclosure No	1	-0.2583	0.8097	-1.8453	1.3286	0.10	0.7497
Scale	1	2.1091	0.2563	1.6621	2.6764		

Table A23: Log-logistic model

Variables Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept	1	-0.8969	2.7604	-6.3073	4.5134	0.11	0.7452
Gender Female	1	0.8628	0.5531	-0.2213	1.9469	2.43	0.1188
Age	1	-0.0125	0.0292	-0.0697	0.0447	0.18	0.6692
	Re	ligion con	npared wi	th Traditio	nalist		
Christian	1	2.0011	0.9669	0.1061	3.8961	4.28	0.0385
Islam Divorced	1	1 8120 -0.0412	1.0370 1.1606	-0.2106 -2.3160	2.2336	0.00	0.9717
Married	1	0.5654	0.8589	-1.1180	2.2488	0.43	0.5103
Single	1	-0.6831	1.1052	-2.8493	1.4830	0.38	0.5365
Weight	1	0.0433	0.0313	-0.0180	0.1045	1.91	0.1665
-	Re	gimen coi	npared w	ith (CBV/I	VVP)		
AZT/3TC/EFV	1	-0.6806	0.7275	-2.1064	0.7451	0.88	0.3494
AZT/3TC/NVP	1	-1.2478	0.6997	-2.6192	0.1236	3.18	0.0745
	WH	IO Clinica	il Stage co	ompared w	ith IV		
WHO I	1	-0.0190	0.8214	-1.6289	1.5910	0.00	0.9816
II	1	0.3399	0.9144	-1.4523	2.1320	0.14	0.7101
III	1	0.9787	0.6348	-0.2656	2.2229	2.38	0.1232
Disclosure No	1	-0.0713	0.7649	-1.5704	1.4278	0.01	0.9257
Scale	1	1.0840	0.1441	0.8354	1.4065		

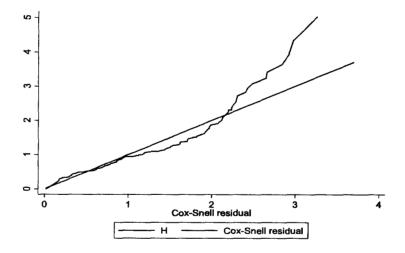


Figure A4: Cox-Snell residual plot for Weibull



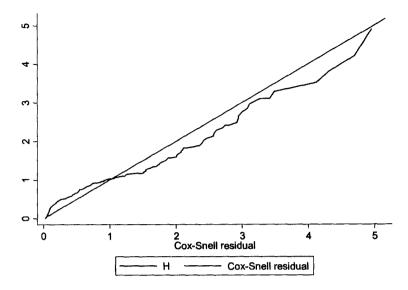


Figure A5: Cox-Snell residual plot for Exponential

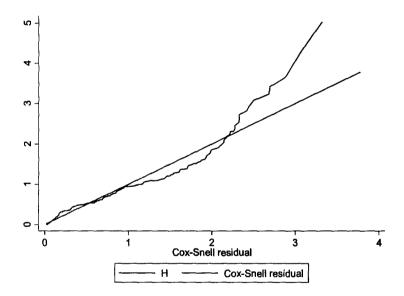


Figure A6: Cox-Snell residual plot for Log-normal

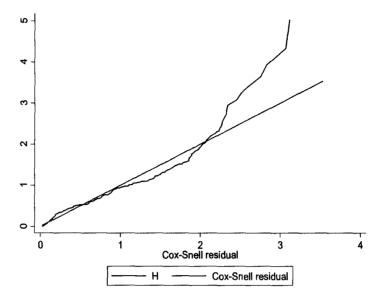


Figure A7: Cox-Snell residual plot for Log-logistic

#### **APPENDIX VI**

#### TABLES FOR AFT MODEL FOR TB PATIENTS

Table A24: Weibull model

Parame	ter Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value
Intercep	t	1	0.7280	1.4408	-2.0959	3.5518	0.26	0.6134
Gender	Female	1	-0.0965	0.4305	-0.9403	0.7472	0.05	0.8226
Age		1	-0.0135	0.0138	-0.0405	0.0135	0.96	0.3261
		Re	ligion com	pared wit	h Traditior	nalist		
	Christian	1	0.6344	0.5834	-0.5091	1.7778	1.18	0.2769
	Islam	1	0.2261	0.6166	-0.9824	1.4346	0.13	0.7139
		Ma	rital status	compared	l with Wid	owed		
	Divorced	1	0.5295	1.0564	-1.5411	2.6001	0.25	0.6162
	Married	1	-0.0040	0.6190	-1.2172	1.2093	0.00	0.9949
	Single	1	0.0197	0.9801	-1.9013	1.9408	0.00	0.9839
Weight		1	0.0437	0.0228	-0.0009	0.0883	3.68	0.0550
		F	Regimen co	ompared w	rith (HRZE	ES)		
	HRZ	1	1.8359	0.6473	0.5671	3.1046	8.04	0.0046
	HRZE	1	1.1560	0.5876	0.0044	2.3076	3.87	0.0491
		T	B type cor	npared wi	th Pulmon	ary		
TB type	Extra Pul	1	0.1470	0.7415	-1.3064	1.6004	0.04	0.8429
Scale		1	0.8895	0.1704	0.6110	1.2950		
Shape		1	1.1242	0.2154	0.7722	1.6365		

Table A25: Exponential model

Variable	es Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept		1	0.6132	1.6119	-2.5461	3.7725	0.14	0.7036
Gender	Female	1	-0.1063	0.4824	-1.0517	0.8391	0.05	0.8256
Age		1	-0.0150	0.0152	-0.0449	0.0148	0.97	0.3236
		Re	ligion con	npared with	h Tradition	alist		
	Christian	1	0.6972	0.6471	-0.5711	1.9655	1.16	0.2813
	Islam	1	0.2454	0.6942	-1.1152	1.6060	0.12	0.7237
		Ma	rital status	compared	with Wid	owed		
	Divorced	1	0.5947	1.1825	-1.7229	2.9123	0.25	0.6150
	Married	1	-0.0051	0.6950	-1.3674	1.3572	0.00	0.9941
	Single	1	0.0330	1.0986	-2.1201	2.1861	0.00	0.9760
Weight		1	0.0500	0.0231	0.0047	0.0952	4.68	0.0305
		R	Regimen co	ompared w	rith (HRZE	ES)		
	HRZ	1	1.9776	0.6825	0.6400	3.3152	8.40	0.0038
	HRZE	1	1.2731	0.6254	0.0474	2.4988	4.14	0.0418
		T	B type con	npared wi	th Pulmon	ary		
Extra Pul	TB	1	0.2115	0.8239	-1.4033	1.8262	0.07	0.7974
Scale		1	1.0000	0.0000	1.0000	1.0000		
Weibull S	Shape	1	1.0000	0.0000	1.0000	1.0000		



Table A26: Log-normal model

Variables	Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept		1	0.3663	1.6095	-2.7884	3.5209	0.05	0.8200
Gender	Female	1	-0.0147	0.4496	-0.8959	0.8665	0.00	0.9739
Age		1	-0.0108	0.0140	-0.0383	0.0166	0.60	0.4388
		Reli	igion com	pared with	Traditiona	alist		
	Christian	1	0.4843	0.6629	-0.8149	1.7835	0.53	0.4650
	Islam	1	0.0798	0.7076	-1.3070	1.4666	0.01	0.9102
		Mari	ital status	compared	with Wide	wed		
	Divorced	1	0.7955	1.1376	-1.4341	3.0250	0.49	0.4844
	Married	1	0.0425	0.6561	-1.2435	1.3284	0.00	0.9484
	Single	1	0.1611	0.9407	-1.6827	2.0049	0.03	0.8640
Weight		1	0.0481	0.0219	0.0051	0.0910	4.80	0.0284
		Re	egimen co	mpared wi	th (HRZE	S)		
	HRZ	1	2.2859	0.7303	0.8545	3.7173	9.80	0.0017
	HRZE	1	1.1857	0.6499	-0.0880	2.4594	3.33	0.0681
		TE	3 type con	pared with	h Pulmona	ıry		
	Extra Pul	1	0.3900	0.8477	-1.2714	2.0515	0.21	0.6454
Scale		_1_	1.6695	0.2888	1.1894	2.3433		

Table A27: Log-logistic model

Variable	es Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value
Intercept		1	0.2043	1.5697	-2.8724	3.2809	0.02	0.8965
Gender	Female	1	-0.0522	0.4407	-0.9160	0.8116	0.01	0.9057
Age		1	-0.0127	0.0137	-0.0395	0.0141	0.87	0.3516
		Re	ligion con	npared wit	h Tradition	alist		
	Christian	1	0.6033	0.6043	-0.5811	1.7877	1.00	0.3181
	Islam	1	0.2169	0.6406	-1.0387	1.4725	0.11	0.7349
		Ma	rital status	compared	l with Wid	owed		
	Divorced	1	0.6362	1.0866	-1.4934	2.7658	0.34	0.5582
	Married	1	0.0354	0.6407	-1.2204	1.2911	0.00	0.9560
	Single	1	0.1056	0.9829	-1.8207	2.0320	0.01	0.9144
Weight		1	0.0471	0.0232	0.0017	0.0925	4.13	0.0420
		F	Regimen co	ompared w	rith (HRZE	ES)		
	HRZ	1	2.0548	0.6891	0.7042	3.4054	8.89	0.0029
	HRZE	1	1.2502	0.6311	0.0132	2.4872	3.92	0.0476
		T	B type con	mpared wi	th Pulmon	ary		
	Extra Pul	1	0.3411	0.8065	-1.2396	1.9218	0.18	0.6723
Scale		1	0.8433	0.1579	0.5843	1.2172		

#### Table A28: Gamma model

Variables Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value			
Intercept	1	1.0491	1.3192	-1.5364	3.6346	0.63	0.4264			
Gender Female	1	-0.1650	0.4197	-0.9877	0.6576	0.15	0.6942			
Age	1	-0.0139	0.0139	-0.0412	0.0133	1.00	0.3164			
	Re	Religion compared with Traditionalist								
Christian	1	0.6472	0.5577	-0.4460	1.7404	1.35	0.2459			
Islam	1	0.2164	0.5881	-0.9364	1.3691	0.14	0.7129			
	Ma	Marital status compared with Widowed								

		0.400-	4 04 - 4	4		<u> </u>	2.552.6
Divorced	1	0.4205	1.0174	-1.5735	2.4145	0.17	0.6794
Married	1	-0.0589	0.5850	-1.2054	1.0877	0.01	0.9199
Single	1	-0.0733	0.9574	-1.9498	1.8031	0.01	0.9389
Weight	1	0.0404	0.0223	-0.0033	0.0841	3.28	0.0703
	F	Regimen co	mpared wi	th (HRZE	S)		
HRZ	1	1.6114	0.6064	0.4228	2.8000	7.06	0.0079
HRZE	1	1.0709	0.5478	-0.0028	2.1446	3.82	0.0506
	T	B type com	pared with	h Pulmona	ry		
Extra Pul	1	-0.0849	0.6783	-1.4143	1.2446	0.02	0.9004
Scale	1	0.3699	0.0730	0.2513	0.5445		

# APPENDIX VII

# TABLES AND FIGURES FOR AFT MODEL FOR HIV/TB

# **CO-INFECTION PATIENTS**

Table A29: Weibull model

Variables Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept	1	-2.5546	1.7573	-5.9987	0.8896	2.11	0.1460
Gender Female	1	1.0524	0.4613	0.1482	1.9566	5.20	0.0225
Age	1	-0.0125	0.0191	-0.0500	0.0250	0.43	0.5139
	Re	eligion con	npared wit	th Tradition	nalist		
Christian	1	0.3891	0.5336	-0.6567	1.4349	0.53	0.4658
Islam	1	0.3943	0.5918	-0.7657	1.5542	0.44	0.5053
	Ma	arital status	s compare	d with Wic	lowed		
Divorced	1	-0.3430	0.6965	-1.7081	1.0222	0.24	0.6224
Married	1	0.2911	0.5612	-0.8089	1.3911	0.27	0.6040
Single	1	3.2286	1.5803	0.1314	6.3259	4.17	0.0410
Weight	1	0.0944	0.0250	0.0455	0.1434	14.31	0.0002
	R	egimen co	mpared wi	ith (CBV/N	√VP)		
AZT/3TC/EFV	1	0.0217	0.5850	-1.1249	1.1683	0.00	0.9704
AZT/3TC/NVP	1	0.3156	0.5140	-0.6918	1.3230	0.38	0.5392
	W]	HO Clinica	al Stage co	mpared w	ith IV		
I	1	-0.3860	0.4859	-1.3383	0.5664	0.63	0.4270
II	1	-0.9368	0.6349	-2.1811	0.3075	2.18	0.1401
III	1	0.5779	0.6914	-0.7772	1.9330	0.70	0.4033
Disclosure No	1	-0.5185	0.4684	-1.4365	0.3995	1.23	0.2683
Scale	1	0.8247	0.1354	0.5977	1.1378		
Shape	1	1.2126	0.1992	0.8789	1.6731		



Table A30: Exponential model

Variables Level	df	β	SE	95%	C.I	χ²	<i>p</i> -value
	1	2 9772	2.0672	-6.9289	1.1744	1.94	0.1640
Intercept	1	-2.8773					
Gender Female	I	1.0791	0.5382	0.0242	2.1339	4.02	0.0450
Age	1	-0.0102	0.0222	-0.0536	0.0333	0.21	0.6467
	Re	ligion com	pared with	n Tradition	alist		
Christian	1	0.4256	0.6310	-0.8111	1.6623	0.45	0.5000
Islam	1	0.3610	0.6939	-0.9991	1.7210	0.27	0.6029
	Ma	rital status	compared	with Wide	owed		
Divorced	1	-0.4007	0.8206	-2.0090	1.2077	0.24	0.6254
Married	1	0.3664	0.6564	-0.9201	1.6528	0.31	0.5767
Single	1	3.4057	1.8051	-0.1321	6.9435	3.56	0.0592
Weight	1	0.0983	0.0295	0.0406	0.1560	11.14	0.0008
•	Re	gimen con	pared wit	h (CBV/N	VP)		
AZT/3TC/EFV	1	0.0971	0.6832	-1.2419	1.4361	0.02	0.8870
AZT/3TC/NVP	1	0.3126	0.6030	-0.8692	1.4945	0.27	0.6041
	WH	IO Clinical	Stage con	mpared wit	th IV		
I	1	-0.3369	0.5866	-1.4867	0.8129	0.33	0.5658
II	1	-0.8940	0.7562	-2.3761	0.5881	1.40	0.2371
III	1	0.4929	0.7771	-1.0301	2.0159	0.40	0.5259
Disclosure No	1	-0.5211	0.5520	-1.6030	0.5608	0.89	0.3452
Scale	1	1.0000	0.0000	1.0000	1.0000		
Shape	1	1.0000	0.0000	1.0000	1.0000		

Table A31: Log-normal model

1 66010 110	71. Log-noi	11141	mouti					
Variables	Level	df	β	SE	95%	C.I	$\chi^2$	<i>p</i> -value
Intercept		1	-1.3353	1.5512	-4.3755	1.7049	0.74	0.3893
Gender	Female	1	0.5629	0.3929	-0.2073	1.3330	2.05	0.1520
Age		1	-0.0146	0.0185	-0.0509	0.0216	0.63	0.4286
		Re	eligion con	npared wit	th Tradition	nalist		
	Christian	1	0.2226	0.5378	-0.8315	1.2767	0.17	0.6789
	Islam	1	0.1725	0.5972	-0.9979	1.3430	0.08	0.7726
		Ma	arital status	compare	d with Wic	lowed		
	Divorced	1	-0.2197	0.7502	-1.6901	1.2507	0.09	0.7696
	Married	1	-0.0215	0.5747	-1.1479	1.1049	0.00	0.9701
	Single	1	1.8603	1.1666	-0.4262	4.1467	2.54	0.1108
Weight		1	0.0786	0.0221	0.0353	0.1219	12.68	0.0004
		R	egimen co	mpared wi	ith (CBV/N	VP)		
AZT/3TC	C/EFV	1	0.1086	0.5697	-1.0080	1.2253	0.04	0.8488
AZT/3TC	C/NVP	1	0.0419	0.4864	-0.9114	0.9953	0.01	0.9313
		$\mathbf{W}$	HO Clinica	al Stage co	mpared w	ith IV		
	I	1	-0.4779	0.5224	-1.5019	0.5460	0.84	0.3603
	II	1	-0.6304	0.6506	-1.9055	0.6447	0.94	0.3326
	III	1	-0.1706	0.5826	-1.3124	0.9713	0.09	0.7697
Disclosur	e No	1	-0.2974	0.4479	-1.1752	0.5804	0.44	0.5066
Scale		1	1.1233	0.1652	0.8420	1.4985		

Table A32: Log-logistic model

Variables	Level	df	β	SE	95% (	C.I	$\chi^2$	<i>p</i> -value
Intercept		1	-1.7967	1.9366	-5.5923	1.9989	0.86	0.3535
Gender	Female	1	0.7145	0.4277	-0.1238	1.5528	2.79	0.0948
Age		1	-0.0152	0.0209	-0.0561	0.0257	0.53	0.4670
_		Re	ligion con	pared with	Traditiona	list		
	Christian	1	0.2826	0.5473	-0.7901	1.3553	0.27	0.6056
	Islam	1	0.2154	0.6386	-1.0363	1.4672	0.11	0.7359
		Ma	rital status	compared	with Wido	wed		
	Divorced	1	-0.2912	0.7119	-1.6866	1.1042	0.17	0.6825
	Married	1	0.0181	0.5908	-1.1399	1.1761	0.00	0.9755
	Single	1	2.0319	1.5282	-0.9633	5.0272	1.77	0.1837
Weight		1	0.0852	0.0233	0.0394 0	.1309	13.33	0.0003
		Re	egimen cor	npared wit	h (CBV/NV	/P)		
AZT/3TC/	EFV	1	0.2085	0.6279	-1.0222	1.4391	0.11	0.7399
AZT/3TC/	NVP	1	0.0917	0.5392	-0.9651	1.1485	0.03	0.8650
		WH	HO Clinica	il Stage cor	npared with	ı IV		
	I	1	-0.4213	0.5192	-1.4390	0.5964	0.66	0.4171
	II	1	-0.7207	0.6408	-1.9767	0.5352	1.27	0.2607
	III	1	-0.2023	-1.5395	1.1350	0.09	0.77	0.6823
Disclosure	No	1	-0.3472	0.4975	-1.3223	0.6280	0.49	0.4853
Scale		1	0.6508	0.1061	0.4728	0.8958		

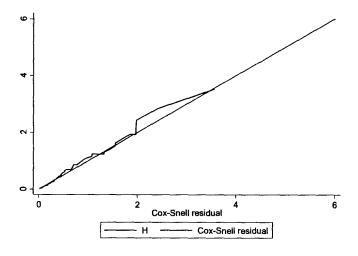


Figure A8: Cox-Snell residual plot for Weibull



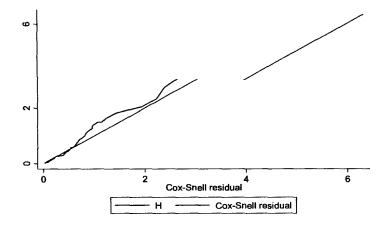


Figure A9: Cox-Snell residual plot for Exponential

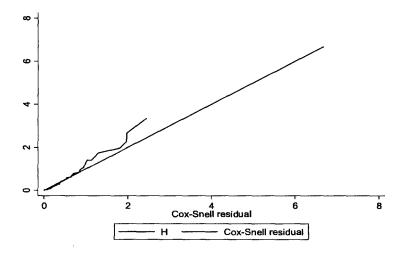


Figure A10: Cox-Snell residual plot for Log-normal



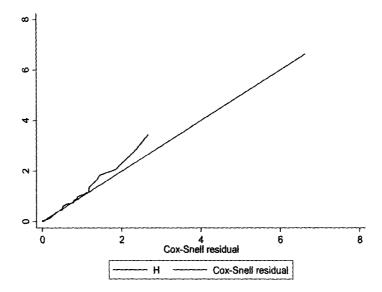


Figure A11: Cox-Snell residual plot for Log-logistic